Bayesian hierarchical models: convexity, sparsity and model reduction

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Problem statement

Want to reconstruct $x \in \mathbb{R}^n$ from few indirect, noisy observations. In the case of a linear observation model

$$b = Ax + e$$
, $A \in \mathbb{R}^{m \times n}$, $m \ll n$.

Assume that

- additive Gaussian noise e; where $E \sim \mathcal{N}(0, I_m)$
- x is believed to be sparse, i.e.,

 $\|x\|_0 \ll n.$

• or to admit a sparse representation

$$x=\mathsf{L} z, \quad \|z\|_0\ll n.$$

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Sparsity Considerations

Sparsity means a signal with a *sparse representation*

- The sparse vector in that case contains the coefficients of a suitable representation, for example
- Wavelet basis
- Fourier basis
- First order differencing matrix for piecewise constant signals in terms of their increments

The conditionally Gaussian random variable is the presumably sparse coefficient vector.

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Sparsity promotion via hierarchical model

• Conditionally Gaussian prior for sparse object

$$X \sim \mathcal{N}(0, \mathsf{D}_{\theta}), \quad \mathsf{D}_{\theta} = \operatorname{diag}(\theta_{1}, \dots, \theta_{n}),$$
 $\pi_{x \mid \theta}(x \mid \theta) = \frac{1}{(2\pi)^{n/2} \sqrt{\theta_{1} \cdots \theta_{n}}} \exp\left(-\frac{1}{2} \sum_{j=1}^{n} \frac{x_{j}^{2}}{\theta_{j}}\right)$

• Mutually independent unknown prior variances $\theta_j > 0$ follow generalized gamma distributions,

$$\Theta_j \sim \text{GenGamma}(r, \vartheta_j, \beta), \quad \pi_{\Theta_j}(\theta_j) = \frac{1}{\Gamma(\beta)\vartheta_j} \left(\frac{\theta_j}{\vartheta_j}\right)^{r\beta-1} \exp\left(-\frac{\theta_j}{\vartheta_j}\right)^r$$

Posterior density

$$\pi_{X,\Theta|B}(x,\theta) \propto \exp\left(-\frac{1}{2}\|b - Ax\|^2 - \frac{1}{2}\sum_{j=1}^n \frac{x_j^2}{\theta_j} + \eta \sum_{j=1}^n \log\frac{\theta_j}{\vartheta_j} - \sum_{j=1}^n \left(\frac{\theta_j}{\vartheta_j}\right)^r\right)$$

where
$$\eta = r\beta - 3/2 > 0$$

Iterated Alternating Sequential (IAS) algorithm

To compute x_{MAP} we minimize the Gibbs energy

$$\mathscr{E}(x;\theta) = \underbrace{\frac{1}{2} \|b - Ax\|^2}_{(\mathscr{P}(x;\theta))} + \underbrace{\sum_{j=1}^n \frac{x_j^2}{2\theta_j}}_{(\mathscr{P}(x;\theta))} - \sum_{j=1}^n \left(\eta \log \frac{\theta_j}{\vartheta_j} - \left(\frac{\theta_j}{\vartheta_j}\right)^r\right)}_{(\mathscr{P}(x;\theta))} \tag{1}$$

Given the initial value $\theta^0 = \vartheta$, $x^0 = 0$, and k = 0, iterate until convergence: (a) Update $x^k \to x^{k+1}$ by minimizing $\mathscr{E}(x \mid \theta^k)$; (b) Update $\theta^k \to \theta^{k+1}$ by minimizing $\mathscr{E}(\theta \mid x^{k+1})$.

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IAS algorithm for Generalized Gamma hyperpriors

• Given
$$\theta$$
, $x_{k+1} = \operatorname{argmin} \left\{ \|b - Ax\|^2 + \|D_{\theta}^{-1/2}x\|^2 \right\}$ solves

$$\begin{bmatrix} A \\ D_{\theta}^{-1/2} \end{bmatrix} x = \begin{bmatrix} b \\ 0 \end{bmatrix}$$

in the least squares sense.

② The update of θ is componentwise. From the first order optimality condition θ_j must satisfy

$$-\frac{1}{2}\frac{x_j^2}{\theta_j^2} - \left(r\beta - \frac{2}{3}\right)\frac{1}{\theta_j} + r\frac{\theta_j^{r-1}}{\vartheta_j^r} = 0, \ x_j = x_j^{t+1}.$$

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Convexity and Convergence: r = 1

For the gamma hyperprior (r = 1):

- $\bullet\,$ The Gibbs energy functional $\mathscr E$ is strictly convex and has a unique minimizer
- In exact arithmetic, the IAS algorithm converges to the global minimizer
- For η > 0 small, the Gibbs energy (1) is approximately equal to the penalized least squares functional with a weighted l₁-penalty.

Theorem

For a gamma hyperprior, the exact IAS algorithm converges to the unique minimizer $(\hat{x}, \hat{\theta})$ of the Gibbs energy functional. Moreover, the minimizer $(\hat{x}, \hat{\theta})$ satisfies the fixed point condition

$$\widehat{x} = \operatorname{argmin} \left\{ \mathscr{E} \left(x \mid F(x) \right) \right\}, \quad \widehat{\theta} = F(\widehat{x}),$$

where F is the map with jth component f_j .¹.

¹Calvetti D, Pascarella A, Pitolli F, Somersalo E, Vantaggi B (2015) A hierarchical Krylov–Bayes iterative inverse solver for MEG with physiological preconditioning. Inverse Problems 31:125005

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Scale parameter and sparsity: r = 1

Under the assumptions of our hierarchical Bayesian model we have shown that

• The exact IAS iteration converges to the global minimizer of the functional

$$\mathscr{L}_{\eta}(x) = \mathscr{E}(x, f(x))$$

and, for small $\eta > 0$

$$\mathscr{L}_{\eta}(x) = \mathscr{L}_{1}(x) + \underbrace{\eta g(x, \eta)}_{\to 0 \text{ as } \eta \to 0},$$

where

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$$\mathscr{L}_{1}(x) = \frac{1}{2} \|b - Ax\|^{2} + \sqrt{2} \sum_{j=1}^{n} \frac{|x_{j}|}{\sqrt{\vartheta_{j}}}.$$

and the sum extends only over the support of x,

$$S = \operatorname{supp}(x) = \{j \mid x_j \neq 0\}.$$

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ℓ_2 Stable Recovery: r = 1

$$\underbrace{x_{\eta} = \operatorname{argmin} \left\{ \mathscr{L}_{\eta}(x) \right\}}_{= IAS \ solution} \qquad \underbrace{x_{1} = \operatorname{argmin} \left\{ \mathscr{L}_{1}(x) \right\}}_{= \ell_{1} \text{ penalized solution}}.$$

- The size of x_η x₁ depends continuously on η. Thus η controls the sparsity of the solution.
- **②** If A is of the kind for which the ℓ_1 -magic works and the data come from a sparse vector², then x_η is close to the underlying sparse solution.
- O The scale parameters ϑ_j play the role of sensitivity weights in inverse problems: Data components may have different sensitivity to different components x_j.

²Candes E, Romberg JK and Tao T(2006): Stable Signal Recovery from Incomplete and Inaccurate Measurements, Comm Pure Appl Math LIX: 1207–1223 · () · () · () · ()

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Sparsity and exchangeability

Assume the underlying signal x is sparse $supp(x) = S \subset \{1, 2, ..., n\}$ and b_0 is the noiseless measurement. Define

$$SNR_{S} = \frac{E\{\|b_{0}\|^{2} | supp(x) = S\}}{E\{\|e\|^{2}\}}, \ e \sim \mathcal{N}(0, \Sigma).$$

Lemma

With our assumptions about X and the noise

$$\mathrm{SNR}_{\mathsf{S}} = \frac{\sum_{j \in \mathsf{I}} \nu(r, \beta) \vartheta_j \|\mathsf{A} e_j\|^2}{\mathrm{tr}\left(\Sigma\right)} + 1, \ \nu(r, \beta) = \frac{\mathsf{\Gamma}(\beta + 1)r}{\mathsf{\Gamma}(\beta r)}.$$

Proof.

$$\textit{E}\left\{ \|\textit{b}_0\|^2 \right\} = \mathrm{Tr}\textit{E}\left\{\textit{b}_0\textit{b}_0^\mathrm{T}\right\} = \mathrm{Tr}\textit{E}\left\{\textit{Axx}^\mathrm{T}\textit{A}^\mathrm{T}\right\} = \mathrm{Tr}\left(\textit{A}\textit{E}\left\{\textit{xx}^\mathrm{T}\right\}\textit{A}^\mathrm{T}\right),$$

and from the generalized gamma hyperprior

$$E\{xx^{\mathrm{T}}\} = E_{\theta}\{E\{xx^{\mathrm{T}} \mid \theta\}\} = E(\operatorname{diag}(\theta)) = \operatorname{diag}(\nu(r,\beta)\vartheta).$$

Scale parameter and sensitivity scaling, in Bayesian way

How should ϑ be chosen?

Theorem

Given an estimate $\overline{\mathrm{SNR}}$ of SNR, if

$$P(||x||_0 = k) = p_k, \quad p_0 = p_n = 0, \quad \sum_{k=1}^n p_k = 1$$

and if

$$\mathrm{SNR}_S = \mathrm{SNR}_{S'}, \quad \forall \; S, S': \mathrm{card}(S) = \mathrm{card}(S'),$$

then

$$artheta_j = rac{\mathsf{C}}{\|\mathsf{A} e_j\|^2}, \quad \mathsf{C} = rac{(\overline{\mathrm{SNR}} - 1)\mathrm{Tr}(\Sigma)}{
u(r, eta)} \sum_{j=1}^n rac{p_k}{k}$$

In the literature $||Ae_j||$ is the *sensitivity* of the data to *j*th component of *x*.

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Sensitivity can make a difference.

Without sensitivity



With sensitivity



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5. Converge rate: r = 1

Sparsity and quadratic convergence: r = 1

For the gamma hyperprior, as η goes to zero the sequence of IAS minimizers remains bounded.

Lemma

There is a constant B > 0 such that

$$||x_{\eta}|| \leq B$$
,

for all η , $0 \le \eta \le \frac{1}{2}$.

Theorem

If the matrix A is such that the minimizer

 $x_1 = \operatorname{argmin}\{F_1(x)\}$

of the ℓ_1 -penalized functional F_1 is unique, then, as $\eta \to 0+$, the minimizers x_η converge to the minimizer x_1 .

Intermezzo: Sparse or compressible?

• Sparsity

If A is a matrix such that the ℓ_1 regularized solution x_1 is sparse, then the solution of the IAS algorithm with $\eta > 0$ small can be made arbitrarily small outside the support of x_1 .

Compressibility

If the components of x_1 are smaller than a threshold outside a set $S \subset \{1, 2, ..., n\}$, the same is true for the IAS solution x_η with a slightly larger threshold when $\eta > 0$ is small enough.

• Bayesian Sparsity is Compressibility

The Bayesian target reconstruction of a sparse signal is a compressible signal.

Convergence of IAS for r = 1

Theorem

In the IAS algorithm, the updates of x converge at least $\hat{\theta}$ -linearly, that is, linearly in the Mahalanobis norm

$$\|x\|_{\widehat{\theta}}^2 = x^{\mathsf{T}}\mathsf{D}_{\widehat{\theta}}^{-1}x$$

evaluated at the MAP estimate. Moreover, if $supp(\hat{x}) \subsetneq \{1, 2, ..., n\}$, the convergence of θ in the complement of the support is quadratic³.

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³D. Calvetti, E.Somersalo and A. Strang. Hierachical Bayesian models and sparsity: ℓ_2 -magic. Inverse Problems 35: 035003.

Generalized gamma hyperpriors and IAS⁴

For the family of generalized gamma hyperpriors for sparse recovery we want to investigate the

- Convexity or lack thereof of Gibbs functional
- Form and behavior of $\boldsymbol{\theta}$ update
- Type of regularization effect on components
- Similarity with classical regularization functionals
- Role of *r* and shape parameter.

Non-dimensionalization:

- WLOG we assume that $\vartheta_j = 1$ or, equivalently,
- scale x_j by $\sqrt{\vartheta_j}$ and θ_j by ϑ_j .

⁴D. Calvetti, M.Pragliola, E. Somersalo and A. Strang. Sparse reconstructions from few noisy data via hierarchical Bayesian models with generalized gamma hyperpriors: convergence, convexity and performance. Manuscript.

The θ update as a function of r

For generalized gamma hyperpriors, the function $\theta_{k+1} = f(x_{k+1})$ is the unique solution of the IVP:

$$\frac{d}{dx}f(x) = \frac{2xf(x)}{2r^2f(x)^{r+1} + x^2}, \quad f(0) = \left(\frac{\eta}{r}\right)^{\frac{1}{r}}, \ x > 0$$

and f(x) = f(-x). Moreover, f is

- Monotonically increasing and unbounded above
- Asymptotically, when |x| is small

$$f(x) \propto \left(\frac{\eta}{r}\right)^{\frac{1}{r}} + \frac{1}{2\eta r}x^2$$

• Asymptotically, when |x| is large

$$f(x) \propto |x|^{p}, \ p = \frac{2}{r+1} \quad r > 0$$

$$f(x) \propto x^{2} \qquad r < 0,$$

with growth linear for r = 1, less than linear r > 1, quadratic r < 1.

Effective local penalty functional

- Shape parameter determines initial value f(0)
- Shape parameter does not affect variance of large |x|

$$\mathscr{P}_j(\mathsf{x}_j \mid \theta_j) = rac{\mathsf{x}_j^2}{2\theta_j} - \eta \log \theta_j + (\theta_j)^r)$$

- For small $|x_j|$: $\mathscr{P}_j(x)$ is quadratic in |x|;
- For large $|x_j|$: $\mathscr{P}_j(x)$ is proportional to
 - $|x_j|^p, p = \frac{2r}{1+r}, r > 0$ • $\log |x_i|, r < 0.$
- When r = 2, p = 4/3.
- When r = 1 p = 1, thus ℓ_1 -like penalty.
- When 0 < r < 1, p < 1 and the penalty strongly enforces sparsity.

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Convexity, Sparsity and Penalization

The Gibbs functional $\mathscr{E}(x,\theta)$ is convex

• for all x, θ if $r \ge 1$ and $\eta > 0$

• for all x,
$$heta < \overline{ heta} = \left(rac{\eta}{r(1-r)}
ight)$$
 if $r < 1$

Convexity region:

- Let $\overline{x} = f^{-1}(\overline{\theta})$. The convexity region is all $x : ||x||_{\infty} < \overline{x}$.
- $\bullet\,$ The radius of the convexity region \overline{x} increases monotonically with $\eta\,$
- $\bullet~\eta$ is proportional to the radius of the convexity region centered at origin

6. Generalized Gamma Hyperpriors



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Support of the signal: the meaning of θ

In light of the Bayesian set up:

- The entries of x with large variance are more likely to contain large values
- The prior variance of x_j is θ_j
- The entries of θ above a threshold identify the support of the signal
- The more sparsity promoting the hyperprior, the more θ greedy the IAS At each IAS iteration, the system learns the support of the signal and uses it to improve the reconstruction.

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IAS with bound constraints

The IAS method can be modified to include bounds on the entries of the solution.

• Assume we believe

$$0 < x_j < H$$

Define

$$G(x) = \left\{ egin{array}{cc} 0, & ext{when } 0 < x \leq H, \ \infty & ext{otherwise}, \end{array}
ight.$$

• Write posterior density with the bound constraints as

$$\pi(x, \theta \mid b) \propto \exp\left(-\mathscr{E}(x, \theta) - \mathcal{G}(x)\right) = \exp\left(-\mathscr{E}_{\mathcal{G}}(x, \theta)\right).$$

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Moreau-Yoshida envelope and box contraints

• Consider the Moreau-Yoshida envelope

$$\Phi^{\lambda}_{G}(x,\theta) = \mathscr{E}(x,\theta) + G^{\lambda}(x),$$

where

$$G^{\lambda}(x) = \min_{u \in \mathbb{R}^n} \left\{ G(u) + \frac{1}{2\lambda} \|x - u\|^2 \right\}, \ \lambda > 0.$$

• The Moreau-Yoshida envelope is differentiable and

$$abla_x \Phi_G^\lambda(x, heta) =
abla_x \mathscr{E}(x, heta) + rac{1}{\lambda} (x - \mathrm{prox}_G^\lambda(x)),$$

where the proximal operator is

$$\begin{aligned} \operatorname{prox}_{G}^{\lambda}(x) &= \operatorname{argmin}_{u \in \mathbb{R}^{n}} \left\{ G(u) + \frac{1}{2\lambda} \|x - u\|^{2} \right\} \\ &= \left\{ \begin{array}{l} x, & \text{if } G(x) = 0, \\ \operatorname{Pz}, & \text{if } G(x) = \infty. \end{array} \right. \end{aligned}$$

and P is the orthogonal projector on the feasible set $[0, H]^n$.

What is the Moreau-Yoshida envelope doing for us?

It has been shown that

- as $\lambda \to 0+$,
- the posterior distribution in terms of the Moreau-Yoshida envelope
- converges to the posterior distribution + positivity constraint.

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IAS with bound constraint

- The inclusion of the bounds does not change ∇_{θ} ,
- The IAS algorithm can be extended for bound constrained problems
- Replace the least squares minimization by the sequential procedure:
- Given the current θ^t:
 - (a) Find $x = x^*$ solving $\nabla_x \mathscr{E}(x, \theta^t) = 0$ in the least squares sense,
 - (b) Update $x^{t+1} = \operatorname{prox}_{G}^{\lambda}(x^{*})$ by projecting x^{*} onto the feasible set.

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Approximate IAS and reduced model

In the case where $\mathsf{A} \in \mathbb{R}^{m imes n}, \ m < n$ at each IAS step, instead of solving

$$\left[\begin{array}{c}\mathsf{A}\\\mathsf{D}_{\theta}^{-1/2}\end{array}\right]\mathsf{x} = \left[\begin{array}{c}\mathsf{b}\\\mathsf{0}\end{array}\right]$$

solve approximately

$$\mathsf{AD}_{\theta}^{1/2}w = b, \quad x = \mathsf{D}_{\theta}^{1/2}w$$

with the CGLS methods equipped with stopping rule.

- Each CGLS iteration requires only 1 matvec with A and one with A'
- If θ_j is small, the corresponding column of $AD_{\theta}^{1/2}$ is almost deflated
- Equivalently, the corresponding solution entry is made smaller
- The more sparsity promoting the prior, the fewer the large θ_j

Three computed examples

- Example 1: Deconvolution of one dimensional staircase signal blurred with Airy kernel⁵. Exact and CGLS-AS
- Example 2: Reconstruction of two dimensional nearly black object recovery from blurred, noisy data (Gaussian blur). Exact and CGLS-IAS
- Example 3: Limited angle computed tomography problem. CGLS-IAS only.

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r = 1, r = 0.5 and r = -0.5 with 1% noise



Starry night: r = 1



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Limited angle tomography



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r = 1

Horizontal and vertical profiles, and CGLS steps



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