# Bayesian Hospital Mortality Rate Estimation and Standardization for Public Reporting 

Edward I. George<br>Veronika Ročková, Paul R. Rosenbaum, Ville A. Satopää and Jeffrey H. Silber

The Wharton School, University of Pennsylvania edgeorge@upenn.edu

Workshop on New and Evolving Roles of Shrinkage in Large-Scale Prediction and Inference Banff International Research Station, Canada

April 11, 2019

## Prologue - Is shrinkage estimation really safe?

■ Observe $x_{i} \mid \mu_{i} \sim N\left(\mu_{i}, v_{i}\right), i=1, \ldots, p$ independently.

- Hierarchical model: $\mu_{1}, \ldots, \mu_{p} \stackrel{\text { iid }}{\sim} N(\mu, v)$.
- $E\left[\mu_{j} \mid\right.$ data $]=\frac{v}{v_{i}+v} x_{i}+\frac{v_{i}}{v_{i}+v} \mu$

■ eB shrinkage estimator: $\hat{\mu}_{j}=\frac{v_{\mu}}{v_{i}+\hat{v}} x_{i}+\frac{v_{i}}{v_{i}+\hat{v}} \bar{x}$

- $\hat{\mu}_{j}$ strongly shrinks $x_{i}$ towards $\bar{x}$ when $v_{i} \gg \hat{v}$
- Justifications for this estimation rely critically on the prior assumption that
$\mu_{1}, \ldots, \mu_{p}$ all have the same mean $\mu$ and variance $v$ !


## Some History - The National Halothane Study (1969)

■ 856,500 surgeries under anesthesia at 34 Hospitals from 1959-1962

- 16,840 deaths within 6 weeks of surgery (about $2 \%$ )
- Compared the effect of general anesthetic agents, especially halothane, on postoperative mortality
- Reported indirectly and directly standardized mortality rates

■ Major differences between hospitals were observed, even after standardization

## THE NATIONAL HALOTHANE STUDY

A STUDY OF THE POSSIBLE ASSOCIATION BETWEEN HALOTHANE ANESTHESIA AND POSTOPERATIVE HEPATIC NECROSIS

Report of
The Subcommittee on the National Halothane Study, of the Committee on Anesthesia, Division of Medical Sciences, National Academy of Sciences-National Research Council

Washington, D.C.

Edited by
John P. Bunker, M. D., William H. Forrest, Jr., M.D. Frederick Mosteller, Ph.D., and Leroy D. Vandam, M.D.

Supported by
Department of Health, Education, and Welfare
Public Health Service, National Institute of General Medical Sciences Contract PH43-63-65

```
MgStre
APR 251969
HD`\mp@code{y }
```

National Institutes of Health
National Institute of General Medical Sciences
Bethesda, Md.

National Alalothane Study: a Study of the Possible Association Between Halothane Anesthesia and Postoperative Hepatic Necrosis; Report. Edited by John P.I http://www.nap.edu/catalog.php?record_id=19006
.N27
C. 1

## MEMBERS OF SUBCOMMITTEE

John P. Bunker, M.D., Chairman
Charles G. Child III, M.D.
Charles S. Davidson, M.D.
Edward A. Gall, M.D.
Gerald Klatskin, M.D.
Leonard Laster, M.D.
Lincoin E. Moses, Ph. D.
Edward A. Gall, M.D., Chairman
Archie H. Baggenstoss, M.D.
I. Nathan Dubin, M.D.

Yvonne M.M. Bishop, Ph. D.
Byron W. Brown, Ph. D.
W. Morven Gentleman, Ph. D.
John P. Gilbert, Ph. D.
Lincoln E. Moses, Ph. D.

## PANEL OF PATHOLOGISTS

Frederick Mosteller, Ph. D.
Shihhsun Ngai, M.D.
Leroy D. Vandam, M.D.
Staff:
William H. Forrest, J., M.D.
Sam F. Seeley, M.D.
John P. Gilbert, Ph. D.

Paul R. Glunz, M.D.
Hans Popper, M.D.
Hans F. Smetana, M.D.

## PANEL OF STATISTICIANS

Frederick Mosteller, Ph. D.
John W. Tukey, Ph. D.
Staff:
Jerry Halpern, M.S.
Lawrence G. Tesler, A.M.

CONSULTANTS AND ASSOCIATES
Clinical Pharmacology:
J. Weldon Bellville, M.D.
Clinical Analyses:
Bernard M. Babior, M.D.
William E. Dozier, M.D.
Pathology:
Charles W. Blumenfeld, M.D.
Beatrice W. Ishak, M.D.
Kamal K. Ishak, M.D.

Project Officers, National Institute of General
Medical Sciences:
Ruth K. Beecroft, M.D.
Carl R. Brewer, Ph. D.
Louis P. Hellman, M.A.

Staff, Division of Medical Sciences,
National Academy of Sciences-
National Research Council
Gilbert W. Beebe, Ph. D.
R. Keith Cannan, Ph. D.

The manuscript was prepared largely at Stanford and at Harvard, with the assistance of the following staff: Stanford Department of Anesthesis: Adena Goodart, Lee Amideo, and Jacqueline Hardy; Stanford Department of Statistics: Ginnie Currey and Margaret Cline; Harvard Department of Statistics: Monica Anvoner, Linda Falcone, Lynn Holtz, Nancy Larson, Holly Lasewicz, Louise Rothman, and Patricia Scott; Harvard Department of Anesthesia: Helen T. Gallahue; and Bell Telephone Laboratories: Elizabeth L. La Jeunesse.

Our thanks are extended to Princeton University, particularly in connection with research sponsored by the Army Research Office (Durham), and Bell Telephone Laboratories for the participation of W. Morven Gentleman and of John W. Tukey; to the Center for Advanced Study in the Behavioral Sciences at Stanford California, and to the University of North Carolina at Chapel Hill, especially through the National Institute of Mental Health's grant MH-10006 to the Psychometric Laboratory, and to the Harvard Computing Center for the services of John P. Gilbert; to Harvard University and the National Science Foundation for facilitating, through grant GS-341, the participation of Frederick Mosteller and Cleo Youtz; to the Harvard School of Public Health for facilitating, through biostatistics training grant 6893-2, the participation of Yvonne M. M. Bishop; and to the Stanford Department of Statistics for facilitating, through biostatistics training grant GM-00025, the participation of Jerry Halpern.

Among our statistical colleagues, Raj Bahadur, James R. Boen, K. A. Brownlee, W. G. Cochran, Arthur P. Dempster, Leo A. Goodman, William Kruskal, Paul Meier, Rupert G. Miller, Jt., Charles Stein, and David L. Wallace have been generous with their advice.

John P. Bunker

## Chairman

Subcommittee on the National
Halothane Study

## Public Reporting of Hospital Mortality Rates Today

■ Medicare's web based "Hospital Compare":
To provide the public "with information on how well the hospitals in your area care for all their adult patients with certain medical conditions" such as heart attacks.
(U.S. Department of Health and Human Services, 2007)

■ Available at http://www.hospitalcompare.hhs.gov

## A Typical Medicare Mortality Rate Report

## Death rate for heart attack patients

Why is this important?

U.S. National Death rate for heart attack patients $=\mathbf{1 5 . 5} \%$

Figure: Comparing heart attack mortality rates with Hospital Compare

## Medicare Public Reporting

- Hospital Compare 2008: Out of 4311 hospitals, 4302 of them (99.8\%) are "no different than U.S. National rate" and zero hospitals are "worse than U.S. National rate".

■ How did Hospital Compare reach these conclusions?

- The smaller the hospital volume, the more its mortality rate estimate is "shrunk" to the overall mean.
- Medicare's justification: Estimates for small volume hospitals rely on the pooled data of all hospitals: "this pooling affords borrowing of statistical strength that provides more confidence in the results."
- Hospital Compare's approach is being copied as the "Gold Standard" for general performance comparisons.
- UH-OH! Hospital Compare's estimates contradict the conventional wisdom that mortality rates are higher at low volume hospitals!!!


## Hospital Compare's Reported Mortality Rates by Volume



Figure: Observed Hospital Rates


Figure: Reported Hospital Rates

Two Major Steps:
1 Hospital Compare begins with a log linear random effects model to predict hospital mortality rates.
2 These rates are then standardized (indirectly) to adjust for patient case-mix differences.

## Administrative Data from Medicare Billing Records

- Medicare data on AMI (Acute Myocardial Infarction) cases from July 1, 2009 to December 31, 2011.
- 377, 615 AMI patients admitted to 4,289 hospitals.
- 56, 567 deaths within 30 days of admission (about $15 \%$ )
- 27 patient characteristics (e.g. age, heart failure, hypertension etc)
- 4 hospital characteristics (volume, resident-to-bed ratio, nurse-to-bed ratio, PCI)
- Training: first 2 years. Validation: remaining 6 months.


## Hospital Compare's Random Effects Model

$$
\log \left(\frac{p_{h j}}{1-p_{h j}}\right)=\alpha_{h}+\boldsymbol{x}_{h j}^{\prime} \boldsymbol{\beta}
$$

where

- $p_{h j}=P\left(Y_{h j}=1\right): 30$-day mortality rate at hospital $h$ for patient $j$, $\left(h=1,2, \ldots, H\right.$ and $\left.j=1,2, \ldots, n_{h}\right)$.
- $\alpha_{h} \sim N\left(\mu, \sigma^{2}\right)$ : hospital random effects.
- $\boldsymbol{x}_{h j}^{\prime} \boldsymbol{\beta}$ : patient fixed effects (based on patient characteristics $\boldsymbol{x}_{h j}$ ).
- Fit using PROC GLIMMIX in SAS, Krumholz et al. (2006ab).


## Hospital Compare Model Estimates



Figure: $P_{h}$ Hospital Mortality Rates


Figure: $\alpha_{h}$ Hospital Effects

- $P_{h}=\frac{1}{n_{h}} \sum_{j=1}^{n_{h}} p_{h j}$ have shrunk the raw observed mortality rates.
- Strong shrinkage of the $\alpha_{h}$ 's for small volume hospitals.
- Under this model, mortality rate $P_{h}$ variation at small volume hospitals driven primarily by patient case-mix differences.


## Shouldn't Hospital Characteristics be in the Model?

■ Hospital effects have been modeled as $\alpha_{h} \sim N\left(\mu, \sigma^{2}\right)$, completely random with the same mean and variance!

- This assumption is leading to the strong shrinkage of the $\alpha_{h}$ 's for small volume hospitals!

■ Available hospital characteristics have been left out!
■ Proponents of the Hospital Compare approach argue that including hospital characteristics in the model would be "unfair" to the hospitals.

■ But isn't excluding hospital characteristics "unfair" to the people seeking accurate information?

Will Adding Hospital Characteristics Make a Difference?

## Suppose we elaborate $\alpha_{h} \sim N\left(\mu, \sigma^{2}\right)$ to $\alpha_{h} \sim N\left(\mu_{h}, \sigma_{h}^{2}\right)$

- $\mu_{h}$ and $\sigma_{h}^{2}$ can now be functions of hospital characteristics!

| Model | $\mu_{h}$ | $\sigma_{h}^{2}$ |
| :--- | :---: | :---: |
| $(\mathrm{C}, \mathrm{C})$ | Constant | Constant |
| $(\mathrm{L}, \mathrm{C})$ | Linear(log vol $\left.{ }_{h}\right)$ | Constant |
| $(\mathrm{S}, \mathrm{L})$ | Spline $\left(\log\right.$ vol $\left._{h}\right)$ | Log-Linear $\left(\right.$ vol $\left.{ }_{h}\right)$ |
| $(\mathrm{SL}, \mathrm{L})$ | Spline $\left(\log\right.$ vol $\left.\left._{h}\right)+{\text { Linear }\left(\text { ptca }_{h}, \text { ntbr }_{h}, \text { rtbr }\right.}_{h}\right)$ | Log-Linear $\left(\right.$ vol $\left._{h}\right)$ |
| $(\mathrm{SLI}, \mathrm{L})$ | $(\mathrm{SL}, \mathrm{L})+\left(\operatorname{log~vol}_{h} \times \operatorname{age}_{h j}\right)$ Interaction |  |

- (C,C) is equivalent to Hospital Compare.
- Each subsequent model nests the previous one.
- Fully Bayesian implementations with non-influential, vague priors.
- MCMC calculations via posterior augmentation with Pólya-Gamma latent variables (Polson, Scott, and Windle 2013).


## Emancipating the Means and Variances with Volume



Figure: Posterior means of $\alpha_{h}$ vs vol $_{h}$

- Dramatic improvements over the ( $\mathrm{C}, \mathrm{C}$ ) model.
- Data speak clearly because simpler models are nested.
- Higher mortality rates at low volume hospitals.


## Adding More Hospital Characteristics and an Interaction


: (SL,L)

(SLI,L)

Figure: Posterior means of $\alpha_{h}{\text { vs } \text { vol }_{h}}$

■ Refinements continue to support higher mortality rates at low volume hospitals

## Have Predictions Really Improved?

## Model Comparisons via Predictive Bayes Factors

- In-sample Bayes Factors are unreliable with vague priors.
- Instead, use out-of-sample Predictive Bayes factors versus (C,C)

$$
B F_{\mathcal{M}_{i} / \mathcal{M}_{c c}}=\frac{\mathbb{P}\left(y_{\text {out }} \mid y_{i n}, \mathcal{M}_{i}\right)}{\mathbb{P}\left(y_{\text {out }} \mid y_{\text {in }}, \mathcal{M}_{c c}\right)}
$$

| Model | $(\mathrm{L}, \mathrm{C})$ | $(\mathrm{S}, \mathrm{L})$ | $(\mathrm{SL}, \mathrm{L})$ | $(\mathrm{SLI}, \mathrm{L})$ |
| :---: | :---: | :---: | :---: | :---: |
| $\log B F$ | 27.54 | 32.13 | 35.46 | 37.96 |

- Vast successive improvements. (SLI,L) clearly best.


## Matched Sample Comparisons of Model Predictions

- Predictive BF's gauge overall model fit.
- But what about the calibration of model predictions with future mortality rates on particular segments of patients?

■ For this purpose, we compared each model's predictions to out-of-sample mortality rates on two sets of patients:

- LV: Patients at low-volume hospitals (bottom 20\%)
- HV: Matched patients at high-volume hospitals (top 20\%)
- Controlling patient risk characteristics through matching provides a clearer comparison of predicted mortality rates between low- and high-volume hospitals.


## Matching Strategy

- Five HV patients are matched to each LV patient.
- Matching is based on minimizing weighted distance between patient characteristics, propensity scores and expected mortalities.
- An example of patient distances

|  | HV Patients |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: |
| LV Patients | 1 | 2 | 3 | 4 | 5 |
| a | 1176.30 | 1371.56 | 482.97 | 399.51 | 380.02 |
| b | 1190.85 | 1389.88 | 498.10 | 427.68 | 394.97 |
| c | 816.24 | 1017.29 | 122.94 | 63.94 | 25.97 |
| d | 1120.22 | 1330.56 | 437.57 | 359.39 | 328.57 |

■ Note that patient c is likely to be matched to patient 5.

## Out-of-Sample Comparisons

|  | Low Volume | High Volume <br> Matched | High Volume <br> All |
| :--- | :---: | :---: | :---: |
| Observed Mortality | 28.3 | 19.8 | 12.4 |
| (C,C) | 23.1 | 21.6 | 12.7 |
| (SLI,L) | 29.6 | 21.0 | 12.4 |

Table: Out-of-sample predicted mortality compared against observed mortality in the matched study of low and high volume hospitals.

At low-volume hospitals, (C,C) is poorly calibrated. It predicted $23.1 \%$ when the actual observed was $28.3 \%$.

# Indirect and Direct Standardization 

## Standardizing Mortality Rates for Public Reporting

- Mortality rate $P_{h}$ at hospital $h$ influenced by patient case-mix.

■ Standardize $P_{h}$ to eliminate this effect of case-mix variation.

- Two approaches:
- Indirect standardization (used by Hospital Compare)
- Direct standardization
- Both approaches make use of the fact that the mortality rate for patient $\boldsymbol{x}_{h j}$ at any hospital $h^{*}$ can be obtained via

$$
p_{h^{*}}\left(\boldsymbol{x}_{h j}\right)=\operatorname{logit}^{-1}\left(\alpha_{h^{*}}+\boldsymbol{x}_{h j}^{\prime} \boldsymbol{\beta}\right) .
$$

■ Note that $p_{h^{*}}\left(\boldsymbol{x}_{h j}\right)$ is a counterfactual unless $h^{*}=h$.

## Indirect Standardization

$$
P_{h}^{I S}=\left(P_{h} / E_{h}\right) \times \bar{y}
$$

where

- $E_{h}=\frac{1}{n_{h}} \sum_{j=1}^{n_{h}}\left[\frac{1}{H} \sum_{h^{*}=1}^{H} p_{h^{*}}\left(\boldsymbol{x}_{h j}\right)\right]$
- $E_{h}$ : Average mortality rate of hospital $h$ patients had they been treated at all $H$ hospitals.
- $\bar{y}$ : national average mortally rate $(\approx 15 \%)$.

Some drawbacks:

- lacks probabilistic justification.
- fails to eliminate case-mix variation effects, except for (C,C).
- systematically underestimates actual hospital mortality rates.


Figure: Indirectly Standardized Mortality Rates $P_{h}^{I S}$ vs. vol ${ }_{h}$.

## Direct Standardization

$$
P_{h}^{D S}=\frac{1}{N} \sum_{h^{*}=1}^{H} \sum_{j=1}^{n_{h^{*}}} p_{h}\left(x_{h^{*} j}\right), \quad N=\sum_{h^{*}=1}^{H} n_{h^{*}}
$$

$P_{h}^{D S}$ : Average mortality rate of all $N$ patients had they been treated at hospital $h$.

Benefits of this approach

+ easier to understand.
+ an interpretable, almost linear scaling of $\alpha_{h}$.
+ eliminates the effect of case-mix variation.
+ is correctly calibrated to actual mortality rates.

: (C,C)

: (L,C)

: (SLI,L)

Figure: Directly Standardized Mortality Rates $P_{h}^{D S}$ vs. vol $_{h}$.
■ red horizontal line - average mortality rate over patients
■ blue horizontal line - average mortality rate over hospitals

## Mortality Rate Uncertainty Quantification

- The variation of the $P_{h}^{D S}$ posterior mean estimates is smaller at the low volume hospitals.
■ However, the posterior mean variation should not be confused with posterior uncertainty of the estimates which is conveyed by the full posterior distribution of the $P_{h}^{D S}$ values.


Figure: $P_{h}^{D S}$ posterior uncertainty at 10 hospitals of varying volume.

## Hospital Classification by Mortality Rates

- The credibility intervals for $P_{h}^{D S}$ can be used to classify hospitals into Low, Average and High mortality according to whether its $95 \%$ interval is entirely below, intersects or is entirely above the overall average morality rate of $15 \%$.

|  | All Hospitals |  |  | Lower Volume Quartile |  |  | Upper Volume Quartile |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\begin{gathered} \text { Counts } \\ (\%) \\ \hline \end{gathered}$ | Low | Average | High | Low | Average | High | Low | Average | High |
| (C,C) | $\begin{gathered} \hline \hline 33 \\ (0.752) \end{gathered}$ | $\begin{gathered} \hline 4333 \\ (98.57) \end{gathered}$ | $\begin{gathered} \hline 30 \\ (0.68) \end{gathered}$ | $\begin{gathered} \hline \hline 0 \\ (0.00) \end{gathered}$ | $\begin{gathered} \hline 1116 \\ (100.00) \end{gathered}$ | $\begin{gathered} \hline \hline 0 \\ (0.00) \end{gathered}$ | $\begin{gathered} \hline \hline 32 \\ (2.91) \end{gathered}$ | $\begin{gathered} \hline 1047 \\ (95.27) \end{gathered}$ | $\begin{gathered} 20 \\ (1.82) \end{gathered}$ |
| (SLI,L) | $\begin{gathered} 58 \\ (1.32) \\ \hline \end{gathered}$ | $\begin{gathered} 3310 \\ (75.30) \\ \hline \end{gathered}$ | $\begin{gathered} 1028 \\ (23.38) \\ \hline \end{gathered}$ | $\begin{gathered} 0 \\ (0.00) \\ \hline \end{gathered}$ | $\begin{gathered} 210 \\ (18.82) \\ \hline \end{gathered}$ | $\begin{gathered} 906 \\ (81.18) \\ \hline \end{gathered}$ | $\begin{gathered} 57 \\ (5.19) \end{gathered}$ | $\begin{gathered} 1038 \\ (94.45) \\ \hline \end{gathered}$ | $\begin{gathered} 4 \\ (0.36) \end{gathered}$ |

Table: Hospital Classifications by Low, Average and High Mortality Rates.

## Attribute Effects

■ With $P_{h}^{D S}$ values, meaningful insights into relationships between hospital mortality rates and PTCA, NTBR are RTBR are readily obtained.


$P_{h}^{D S}$ vS NTBR

: $P_{h}^{D S}$ vS RTBR

Figure: $P_{h}^{D S}$ under the ( $\mathrm{SLI}, \mathrm{L}$ ) model.

## Conclusions

- Strong evidence that hospital characteristics and interactions should be included in the model.
- Indirect standardization fails to eliminate the effect of case-mix variation and underestimates actual mortality rates.
- Directly standardized rates should be the new gold standard for public reporting and for further analyses of what influences mortality.


## Dilemma

- Should Medicare publicly report the alarmingly high mortality rates at the low volume hospitals?


## Thank you!

References:
1 Mortality Rate Estimation and Standardization for Public Reporting: Medicare's Hospital Compare.
E. George, V. Ročková, P. Rosenbaum, V. Satopää, J. Silber. Journal of the American Statistical Association 2017.

2 Improving Medicare's Hospital Compare Mortality Model.
J. Silber, V. Satopää, N. Mukherjee, V. Ročková, W. Wong, A. Hill,
O. Even-Shoshan, P. Rosenbaum, E. George. Health Services Research 2016.

