Optimal designs for longitudinal studies with fractional polynomial models

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Joint work with Víctor Casero–Alonso & Weng Kee Wong

Latest Advances Theory & Applications of Design & Analysis of Experiments,

Banff 2017

Reproducible Science

nature human behaviour

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A manifesto for reproducible science

Marcus R. Munafö^{1,2}*, Brian A. Nosek^{3,4}, Dorothy V. M. Bishop⁵, Katherine S. Button⁶, Christopher D. Chambers⁷, Nathalie Percie du Sert⁸, Uri Simonsohn⁹, Eric-Jan Wagenmakers¹⁰, Jennifer J. Ware¹¹ and John P. A. Ioannidis^{12,13,14}

Improving the reliability and efficiency of scientific research will increase the credibility of the published scientific literature and accelerate discovery. Here we argue for the adoption of measures to optimize key elements of the scientific process: methods, reporting and dissemination, reproducibility, evaluation and incentives. There is some evidence from both simulations and empirical studies supporting the likely effectiveness of these measures, but their broad adoption by researchers, institutions, funders and journals will require iterative evaluation and improvement. We discuss the goals of these measures, and how they can be implemented, in the hope that this will facilitate action toward improving the transparency, reproducibility and efficiency of scientific research.

hat proportion of published research is likely to be false? Low sample size, small effect sizes, data dredging (also so so the size small effect sizes, data dredging (also scientists working competitively in silos without combinine their efforts. and so on. max consoire to dramatically increase

The problem

A hallmark of scientific creativity is the ability to see novel and unexpected patterns in data. John Snow's identification of links between cholera and water supply¹⁷, Paul Broca's work on language lateralization¹⁶ and locelvn Bell Burnell's discoverv of pulsars¹⁷ are

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• Claims for a rigorous research methodology.



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- Key measures to optimize the scientific process.

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Demands " (...) the process of describing in full the study design and data collected that underlie the results reported, rather than a curated version of the design, and/or a subset of the data collected".

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Por una investigación de calidad (http://www.elespanol.com/ opinion/tribunas/20170227/197100289_12.html)

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2 Optimal design notation.



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- **2** Optimal design notation.
- **8** D- and I-optimal designs for FP1, FP2 and FP3 models.

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 - 8 Multi-factor models.
- 6 Conclusions.

Robust estimation



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Robust estimation



Maximum Likely Look Estimator (MLLE)

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FP models

Royston & Altman (1994)

$$\phi_2(x; \mathbf{p}) = \alpha_0 + \alpha_1 x^{(p_1)} + \alpha_2 x^{(p_2)}$$



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Fractional Polynomial (FP) models

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$$\phi_m(x; \mathbf{p}) = \alpha_0 + \sum_{j=1}^m \alpha_j H_j(x)$$

• $H_1(x) = x^{(p_1)}$
 $H_j(x) = \begin{cases} x^{(p_j)}, & \text{if } p_j \neq p_{j-1}, \\ H_{j-1}(x) \ln[x], & \text{if } p_j = p_{j-1}, \end{cases}$ for $j = 2, \dots, m$.

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• $x^{(p_j)} = \begin{cases} \ln[x] & \text{if } p_j = 0 \\ x^{p_j} & \text{otherwise} \end{cases}$ (Box-Tidwell transformation)

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Fractional Polynomial (FP) models

$$\phi_{m}(x; \mathbf{p}) = \alpha_{0} + \sum_{j=1}^{m} \alpha_{j} H_{j}(x)$$
• $H_{1}(x) = x^{(p_{1})}$
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• $x^{(p_{j})} = \begin{cases} \ln[x] & \text{if } p_{j} = 0 \\ x^{p_{j}} & \text{otherwise} \end{cases}$ (Box-Tidwell transformation)
• $\mathbf{p} = (p_{1}, \dots, p_{m})$ with $p_{j} \in \mathcal{P} = \{-2, -1, -\frac{1}{2}, 0, \frac{1}{2}, 1, 2, 3\}$
 $(p_{1} \leq \dots \leq p_{m})$
 $x \neq 0(> 0)$

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Design Theory

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- Approximate designs: $\xi = \left\{ \begin{array}{ccc} x_1 & x_2 & \dots & x_k \\ w_1 & w_2 & \dots & w_k \end{array} \right\} \quad x_i \in \chi$
 - ξ is implemented by realizing about nw_i experiments at x_i

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•
$$M(\xi) = \int_{\chi} f(x)f(x)^T \xi(dx)$$

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• Criteria:

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• $\Phi_I(\xi) = \int_S f(x)^T M^{-1}(\xi) f(x) \mu(dx) = tr A M^{-1}(\xi),$

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 μ , user-selected weighting measure over *S*

•
$$T_{21}(\xi) = \min_{\theta_2 \in \Theta_2} \left[\int_{\mathcal{X}} \{\eta(x, \theta_1) - \eta_2(x, \theta_2)\}^2 \xi(dx) \right]$$
.
(assuming θ completely known).

Optimality

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- Equivalence Theorems:
 - $f(x)^T M^{-1}(\xi_D^*) f(x) (m+1) \leq 0$ for all $x \in X$.
 - $f(x)^T M^{-1}(\xi_l^*) A M^{-1}(\xi_l^*) f(x) tr A M^{-1}(\xi_l^*) \le 0$ for all $x \in X$.
 - $\max_{x} \psi(x, \xi_{s}) \leq 0$ for all $x \in X$, where $\psi(x, \xi_{s}) = [f^{T}(x)\theta - f_{1}^{T}(x)\hat{\theta}_{1}]^{2} - \int_{\chi} [f^{T}(x)\theta - f_{1}^{T}(x)\hat{\theta}_{1}]^{2}\xi(dx),$ and $\hat{\theta}_{1} = \arg \min_{\theta_{1}} \int_{\chi} [f^{T}(x)\theta - f_{1}^{T}(x)\theta_{1}]^{2}\xi(dx).$

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• Efficiencies:
$$\left(\frac{|M(\xi)|}{|M(\xi_D^{\star})|}\right)^{\frac{1}{m+1}}$$
, $\frac{\Phi_I(\xi_I^{\star})}{\Phi_I(\xi)}$, $\frac{T_{21}(\xi)}{T_{21}(\xi_T^{\star})}$.

Optimal designs for FP models

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Optimal designs for FP models

•
$$\chi = [\epsilon, a]$$
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Optimal designs for FP models

- $\chi = [\epsilon, a]$,
- Closed-formed formulae,


Optimal designs for FP models

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- $\chi = [\epsilon, a]$,
- Closed-formed formulae,
- A user-friendly applet



http://areaestadistica.uclm.es/oed/index.php/
computer-tools/

D-optimality:

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D-optimality:

 Tchebyshev system (Karlin & Studden, 1966): No non-trivial polynomial in this system has at most n – 1 zeros, counting multiplicity.

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- For FP1, the GET for D-optimality says $c(x) = f^{T}(x)M^{-1}(\xi^{\star})f(x) 2 \leq 0$ for all $x \in [\epsilon, a]$.

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- c(x) is a linear combination of $1, x^p, x^{2p}$.

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- c(x) is a linear combination of $1, x^p, x^{2p}$.
- They form a Tchebyshev system on the interval $[\epsilon, a]$ because

$$\begin{vmatrix} 1 & x_1^p & x_1^{2p} \\ 1 & x_2^p & x_2^{2p} \\ 1 & x_3^p & x_3^{2p} \end{vmatrix} = -(x_1^p - x_2^p)(x_1^p - x_3^p)(x_2^p - x_3^p),$$

has the same sign for any $\epsilon \leq x_1 < x_2 < x_3 \leq a$.

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Then for each 0 ≠ p ∈ P, c(x) has at most 2 zeros, so the D-optimal design is equally supported at 2 points (Pukelsheim, 1993; Fedorov, 1972).

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- Direct calculations show that a design equally supported at the two end-points is D-optimal:

$$c(x) = 4(a^p - x^p)(\epsilon^p - x^p)/(a^p - \epsilon^p)^2 \le 0, p \ne 0.$$

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• A similar argument applies when p = 0.

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I–optimality:

• Components of the sensitivity function: $1, x^p, x^{2p}$ when $p \neq 0$ and $\{1, \ln[x], \ln[x]^2\}$ when p = 0.

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- Components of the sensitivity function: $1, x^{p}, x^{2p}$ when $p \neq 0$ and $\{1, \ln[x], \ln[x]^{2}\}$ when p = 0.
- They form a Tchebyshev system on $[\epsilon, a]$.
- The weights are found by finding the roots of the sensitivity function of the design supported at x = ε and x = a.

$$\xi_D^{\star} = \left\{ \begin{array}{cc} \epsilon & \mathbf{a} \\ 1/2 & 1/2 \end{array} \right\}$$



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$$\xi_D^{\star} = \left\{ \begin{array}{cc} \epsilon & a \\ 1/2 & 1/2 \end{array} \right\} \qquad \qquad \xi_I^{\star} = \left\{ \begin{array}{cc} \epsilon & a \\ w & 1-w \end{array} \right\} \\ \frac{1}{w} = 1 + \sqrt{\frac{(p+1)a^{2p+1} + a(2p+1)[(p+1)\epsilon^{2p} - 2(a\epsilon)^p] - 2p^2\epsilon^{2p+1}}{2p^2a^{2p+1} - (p+1)(2p+1)\epsilon a^{2p} + \epsilon[2(2p+1)(a\epsilon)^p - (p+1)\epsilon^{2p}]}} \right\}$$



SAC

• c(x) has at most 6 components: $1, x^p, x^{2p}, x^q, x^{2q}$ and x^{p+q} .

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- c(x) has at most 6 components: $1, x^p, x^{2p}, x^q, x^{2q}$ and x^{p+q} .
- The Wronskians (Gasull et al., 2012) are positive for any *p* and *q*,

$$\begin{array}{cccccccc} f_1(x) & f_2(x) & \cdots & f_k(x) \\ f_1'(x) & f_2'(x) & \cdots & f_k'(x) \\ \cdots & \cdots & \cdots & \cdots \\ f_1^{(k)}(x) & f_2^{(k)}(x) & \cdots & f_k^{(k)}(x) \end{array},$$

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- Thus, they form a Tchebyshev system.
- There are at most 5 zeroes (counting multiplicities).
- The interior support points have multiplicity two.
- Thus, only three support points are possible: either 1 or 2 interior support points.

• Suppose an equally weighted design supported at $s_1 < s_2 < a$.

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• If $0 \neq p \leq q$,

$$|M(\xi)| = rac{1}{27} \left| egin{array}{ccc} 1 & s_1^p & s_1^q \ 1 & s_2^p & s_2^q \ 1 & a^p & a^q \end{array}
ight|^2 = rac{1}{27} D^2 > 0$$

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• *D* is always either positive or negative for any values of *s*₁, *s*₂ and *a* (Chebyshev system).

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- If $0 \neq p \leq q$,

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- *D* is always either positive or negative for any values of *s*₁, *s*₂ and *a* (Chebyshev system).
- Moreover,

$$\frac{\partial D}{\partial s_1} = q s_1^{q-1} (a^p - s_2^p) - p s_1^{p-1} (a^q - s_2^q) < 0$$

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implies that $\frac{\partial |M(\xi)|}{\partial s_1} = (2/27)D\frac{\partial D}{\partial s_1} < 0.$

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implies that $\frac{\partial |M(\xi)|}{\partial s_1} = (2/27)D\frac{\partial D}{\partial s_1} < 0.$ • Thus, *D* is a decreasing function of s_1 .

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implies that $\frac{\partial |M(\xi)|}{\partial s_1} = (2/27)D\frac{\partial D}{\partial s_1} < 0.$

- Thus, D is a decreasing function of s_1 .
- Consequently, ϵ is a support point of the D-optimal design.

• Last equation holds iff
$$s_1^{q-p} \frac{q(a^p - s_2^p)}{p(a^q - s_2^q)} < q1.$$

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 The interior support point is the unique root of the derivative of the sensitivity function.
... for FP2(p,q)

$$\xi_D^{\star} = \left\{ \begin{array}{cc} \epsilon & s & a \\ 1/3 & 1/3 & 1/3 \end{array} \right\} \quad \dots$$

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... for FP2(p,q)



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$$\xi_I^{\star} = \left\{ \begin{array}{ccc} \epsilon & s & a \\ w_1 & w_2 & 1 - w_1 - w_2 \end{array} \right\}$$



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... FP3(p,q,r)

$$\xi_D^{\star} = \left\{ \begin{array}{ccc} \epsilon & s_1 & s_2 & a \\ 1/4 & 1/4 & 1/4 & 1/4 \end{array} \right\} \quad \xi_I^{\star} = \left\{ \begin{array}{ccc} \epsilon & s_1 & s_2 & a \\ w_1 & w_2 & w_3 & 1 - \sum_i w_i \end{array} \right\}$$



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Model Uncertainty (FP2)



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Order in FP2 models





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Introduced by Atkinson and Fedorov (1975a, b) for discriminating between two rival linear models

$$T_{21}(\xi) = \min_{\theta_2 \in \Theta_2} \left[\int_{\mathcal{X}} \{\eta(x,\theta_1) - \eta_2(x,\theta_2)\}^2 \xi(dx) \right].$$

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KL–optimality for any distribution (López-Fidalgo, Tommasi and Trandafir, 2007):

$$I_{21}(\xi) = \min_{\theta_2 \in \Theta_2} \left\{ \int_{\mathcal{X}} \mathcal{I}(f, f_2, x, \theta_2) \xi(dx) \right\}$$

where $\mathcal{I}(f, f_2, x, \theta_2) = \int f(y, x, \tau) \log \left\{ \frac{f(y, x, \tau)}{f_2(y, x, \theta_2, \tau)} \right\}$ is the Kullback–Leibler (KL) distance.

General KL-opt algorithm

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1 Given a design ξ_s at step s, compute

$$\begin{aligned} \theta_{2,s} &= \arg\min_{\theta_2 \in \Theta_2} \left\{ \int_{\mathcal{X}} \mathcal{I}(f, f_2, x, \theta_2) \xi(dx) \right\} \\ x_s &= \arg\max_{x \in \mathcal{X}} \{ \mathcal{I}(f, f_2, x, \theta_{2,s}) \}. \end{aligned}$$

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$$x_s = \arg\max_{x \in \mathcal{X}} \{ \mathcal{I}(f, f_2, x, \theta_{2,s}) \}.$$

2 Then

$$\begin{split} \xi_{s+1} &= (1 - \alpha_s)\xi_s + \alpha_s\xi_{x_s} \\ (0 \leq \alpha_s \leq 1, \lim_{s \to \infty} \alpha_s = 0, \sum_{s=0}^{\infty} \alpha_s = \infty, \sum_{s=0}^{\infty} \alpha_s^2 < \infty). \end{split}$$

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$$\xi_{s+1} = (1 - \alpha_s)\xi_s + \alpha_s\xi_{x_s}$$

 $(0 \le \alpha_s \le 1, \lim_{s \to \infty} \alpha_s = 0, \sum_{s=0}^{\infty} \alpha_s = \infty, \sum_{s=0}^{\infty} \alpha_s^2 < \infty).$

8 The stopping rule for the algorithm is based on the GET

$$\left[1+\frac{\max_{x\in\mathcal{X}}\psi(x,\xi_s)}{l_{21}(\xi_s)}\right]^{-1} > \delta(=0.999)$$

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Some results

$$\begin{split} f &= \mathsf{FP1}(0); \ f_2 = \mathsf{FP1}(1/2) \quad \left\{ \begin{array}{ccc} 0.01 & 0.153269 & 1 \\ \frac{73}{216} & \frac{1}{2} & \frac{35}{216} \end{array} \right\} \\ f &= \mathsf{FP1}(1/2); \ f_2 = \mathsf{FP1}(0) \quad \left\{ \begin{array}{ccc} 0.01 & 0.152248 & 1 \\ \frac{19}{93} & \frac{1}{2} & \frac{55}{186} \end{array} \right\} \\ f &= \mathsf{FP1}(0); \ f_2 = \mathsf{FP1}(3) \quad \left\{ \begin{array}{ccc} 0.01 & 0.417462 & 1 \\ \frac{89}{192} & \frac{1}{2} & \frac{7}{192} \end{array} \right\} \\ f &= \mathsf{FP1}(-2); \ f_2 = \mathsf{FP1}(3) \quad \left\{ \begin{array}{ccc} 0.01 & 0.148491 & 1 \\ \frac{263}{528} & \frac{1}{2} & \frac{1}{528} \end{array} \right\} \end{split}$$

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Efficiencies for FP1(p)

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	D-eff	D-eff	l-eff	l-eff	T-eff	T-eff
	(0)	(1/2)	(0)	(1/2)	(0, 1/2)	(1/2, 0)
$\xi_D(0, 1/2)$	100	100	80.9	95.6	0.0	0.0
$\xi_{I}(0)$	87.4	87.4	100	91.2	0.0	0.0
$\xi_{I}(1/2)$	97.7	97.7	92.8	100	0.0	0.0
$\xi_T(0, 1/2)$	71.4	66.2	55.7	51.4	100	85.9
$\xi_T(1/2,0)$	69.6	74.5	73.5	75.1	87.0	100

The value of p is between parentheses.

Applications to biomedical studies

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Chitty et al. (1993):

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Chitty et al. (1993):

• Fetal measurements of mandible length for 158 fetuses between 12 and 28 weeks.

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- Fetal measurements of mandible length for 158 fetuses between 12 and 28 weeks.
- The logarithm of the mandible length given the gestational age is approximately homoscedastic and normally distributed.

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- The logarithm of the mandible length given the gestational age is approximately homoscedastic and normally distributed.
- Royston and Altman (1994):
 - Goodness-of-fit of FP1 and FP2 models.
 - The best were FP1(-1) and FP2(-2,1).



Fig. 3. Extrapolated fit for the mandible data (shown on a log-scale) using two models: ——, fractional polynomial $\phi_1(X; -1)$; ------, cubic polynomial

$$\xi_D^{\star} = \left\{ \begin{array}{cc} 12 & 28\\ 1/2 & 1/2 \end{array} \right\} \qquad \xi_I^{\star} = \left\{ \begin{array}{cc} 12 & 28\\ 0.4226 & 0.5774 \end{array} \right\} \\ \mu = U[12, 28] \end{array}$$

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Fig. 3. Extrapolated fit for the mandible data (shown on a log-scale) using two models: ——, fractional polynomial $\phi_1(X; -1)$; ------, cubic polynomial

$$\begin{aligned} \xi_D^{\star} &= \left\{ \begin{array}{cc} 12 & 28 \\ 1/2 & 1/2 \end{array} \right\} & \xi_I^{\star} &= \left\{ \begin{array}{cc} 12 & 28 \\ 0.4226 & 0.5774 \end{array} \right\} \\ \mu &= U[12, 28] \\ \text{I-eff}(\xi_D^{\star}) &= 97.7\% \end{array} \\ \text{D-eff}(\xi_I^{\star}) &= 97.6\% \end{aligned}$$

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Isaacs et al. (1983): Serum immunoglobulin G.

• IgG concentration: Monoclonal gammopathies and immune deficiencies in children between 6 months and 6 years old.

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- 298 independent observations.

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- IgG was skewed and was quite effectively removed by a square root transformation.
- STATA or mfp package from R: Best fitting for FP2(-2,2).
- Clinicians were interested in the IgG levels for children aged between 6 and 7 years old.
- Question was how best predict the IgG levels for this age group.

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Fig. 5. Fits for IgG data: (a) $\phi_2(X; -2, 2)$ (-----), quartic (------); (b) $\phi_2(X; \frac{1}{2}, 1)$ (-----), cubic (------)

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$$\xi_D^{\star} = \begin{cases} 0.5 & 1.7321 & 6\\ 1/3 & 1/3 & 1/3 \end{cases}$$

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Fig. 5. Fits for IgG data: (a) $\phi_2(X; -2, 2)$ (-----), quartic (------); (b) $\phi_2(X; \frac{1}{2}, 1)$ (-----), cubic (------)

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$$\begin{split} \xi_D^{\star} &= \begin{cases} 0.5 & 1.7321 & 6 \\ 1/3 & 1/3 & 1/3 \end{cases} \\ \text{D-eff}(\xi_{implem.}) &= 53.2\% \end{split}$$



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Application 3: longitudinal studies (growth curve with FP)

Advances in Bioinformatics



FIGURE 2: Time-course expression patterns for the 15 significant genes plotted according to the estimated power for transformation and sign of the regression coefficient.

... using appropriate OED theory

• Longitudinal study (gene expression data).

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$$\begin{array}{l} \mathsf{FP1}(\text{-}0.5): \ \xi_D^{\star} = \begin{cases} 0 & 24 \\ 0.5 & 0.5 \end{cases} \ \xi_I^{\star} = \begin{cases} 0 & 24 \\ 0.2280 & 0.7720 \end{cases}, \ U[0, 24] \\ \cdots \\ \mathsf{FP1}(3): \ \xi_D^{\star} = \begin{cases} 0 & 24 \\ 0.5 & 0.5 \end{cases} \ \xi_I^{\star} = \begin{cases} 0 & 24 \\ 0.6726 & 0.3274 \end{cases}, \ U[0, 24] \end{aligned}$$

- Longitudinal study (gene expression data).
- Linear mixed effects model.
- Optimal designs (Prus and Schwabe, 2016)
- Random intercept: Results depend on the dispersion matrix, cov(α_i).

$$\begin{array}{l} \mathsf{FP1}(\text{-}0.5): \ \xi_D^{\star} = \begin{cases} 0 & 24 \\ 0.5 & 0.5 \end{cases} \ \xi_I^{\star} = \begin{cases} 0 & 24 \\ 0.2280 & 0.7720 \end{cases}, \ U[0, 24] \\ \dots \\ \mathsf{FP1}(3): \ \xi_D^{\star} = \begin{cases} 0 & 24 \\ 0.5 & 0.5 \end{cases} \ \xi_I^{\star} = \begin{cases} 0 & 24 \\ 0.6726 & 0.3274 \end{cases}, \ U[0, 24] \\ \mathsf{Obs}: \ \mathsf{For} \ \xi_I^{\star}: \end{array}$$

$$\frac{1}{w} = 1 + \sqrt{\frac{(p+1)a^{2p+1} + a(2p+1)\left[(p+1)\epsilon^{2p} - 2(a\epsilon)^p\right] - 2p^2\epsilon^{2p+1}}{2p^2a^{2p+1} - (p+1)(2p+1)\epsilon a^{2p} + \epsilon\left[2(2p+1)(a\epsilon)^p - (p+1)(a\epsilon)^p\right]}}$$

1 Product type designs (Rafajlowicz and Myszka, 1992).

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2 Multiplicative or additive regression functions.

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1 Product type designs (Rafajlowicz and Myszka, 1992).

- 2 Multiplicative or additive regression functions.
- **3** $\xi_1^D \otimes \ldots \otimes \xi_k^D$ D-optimal (multiplicative or additive).

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- 4 ξ^l₁ ⊗ ... ⊗ ξ^l_k l-optimal under stringent conditions on μ:
 1 Multiplicative model: independent marginals μ₁,..., μ_k,.
 2 Additive model: ∫_{χ_i} f^{*}_i(x_i)μ^{*}_i(dx_i) = 0, for i = 1,..., k.

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computer-tools/
```

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Selected references

Fedorov V.V. (1972), Theory of optimal experiments, Academic Press.



López-Fidalgo J and Tommasi Ch and Trandafir C (2007) An optimal experimental design criterion for discriminating between non-normal models. Journal of the Royal Statistical Society, Series B 69, 231–242,



Royston P. and Altman D.G. (1994), *Regression Using Fractional Polynomials of Continuous Covariates: Parsimonious Parametric Modelling*, Applied Statistics, 43, 429-467.



Royston P. and Sauerbrei W. (2004), A new approach to modelling interactions between treatment and continuous covariates in clinical trials by using fractional polynomials, Statist. Med., 23, 2509–2525.



Prus M. and Schwabe R. (2016) *Optimal designs for the prediction of individual parameters in hierarchical models*, J. R. Statist. Soc. B, 78, 175?-191



Rafajłowicz E. and Myszka W. (1992), When product type experimental design is optimal? Brief survey and new results, Metrika, 39, 321–333,



Wong W.K. (1994), *G*-optimal designs for multi-factor experiments with heteroscedastic errors, Journal of Statistical Planning and Inference, 40, 127–133.