# Pair-copula constructions even more flexible than copulas 

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## Motivation

- While there is a multitude of bivariate copula, the class of multivariate copulae is still quite restricted.
- Hence, if the dependency structures of different pairs of variables in a multivariate problem are very different, not even the copula approach will allow for the construction of an appropriate model.
- In this talk we will describe an extension to the state-of-the-art theory of copulas, modelling multivariate data using a so-called pair-copula construction (PCC).






## Overview

(1) Motivation and background
(2) Pair-copula constructions
(3) How can we estimate and model select PCCs ?
(4) Application: Market risk model for largest Norwegian bank
(5) Recent advances for vines
(6) Summary and outlook

## Copula...

## Theorem (Sklar 1959)

Sklar's theorem states that every multivariate distribution $F$ with marginals $F_{1}\left(x_{1}\right), \ldots, F_{d}\left(x_{d}\right)$ can be written as:

$$
F\left(x_{1}, \ldots, x_{d}\right)=C\left(F_{1}\left(x_{1}\right), \ldots, F_{d}\left(x_{d}\right)\right)
$$

for some $d$-dimensional copula $C$.

Moreover, for an absolutely continuous joint distribution $F$ with strictly increasing continuous marginal distribution functions $F_{1}, \ldots F_{d}$ it holds that

$$
f\left(x_{1}, \ldots, x_{d}\right)=c\left(F_{1}\left(x_{1}\right), \ldots, F_{d}\left(x_{d}\right)\right) \cdot\left[\prod_{i=1}^{d} f_{i}\left(x_{i}\right)\right]
$$

for some $d$-dimensional copula density $c$.

## Pair-copula constructions (I)

- For two random variables $X_{1}$ and $X_{2}$ we have

$$
f\left(x_{1} \mid x_{2}\right)=c_{12}\left(F_{1}\left(x_{1}\right), F_{2}\left(x_{2}\right)\right) \cdot f_{1}\left(x_{1}\right)
$$

- Further, for three random variables $X_{1}, X_{2}$ and $X_{3}$ we have

$$
f\left(x_{1} \mid x_{2}, x_{3}\right)=c_{13 ; 2}\left(F_{1 \mid 2}\left(x_{1} \mid x_{2}\right), F_{3 \mid 2}\left(x_{3} \mid x_{2}\right)\right) \cdot f_{1 \mid 2}\left(x_{1} \mid x_{2}\right)
$$

- It follows that for every $j$ we have

$$
f(x \mid \mathbf{v})=c_{x v_{j} ; \mathbf{v}_{-j}}\left(F\left(x \mid \mathbf{v}_{-j}\right), F\left(v_{j} \mid \mathbf{v}_{-j}\right)\right) \cdot f\left(x \mid \mathbf{v}_{-j}\right)
$$

## Pair-copula constructions (II)

By combining the two results

$$
f\left(x_{1}, \ldots x_{d}\right)=f_{d}\left(x_{d}\right) \cdot f\left(x_{d-1} \mid x_{d}\right) \cdots f\left(x_{1} \mid x_{2}, \ldots x_{d}\right)
$$

and

$$
f(x \mid \mathbf{v})=c_{X v_{j} ; \mathbf{v}_{-j}}\left(F\left(x \mid \mathbf{v}_{-j}\right), F\left(v_{j} \mid \mathbf{v}_{-j}\right)\right) \cdot f\left(x \mid \mathbf{v}_{-j}\right)
$$

we may derive a decomposition of $f\left(x_{1}, \ldots x_{d}\right)$ that only consists of marginal distributions and bivariate copulae.

We denote a such decomposition a pair copula construction (PCC)
Joe (1996) was the first to give a probabilistic construction of multivariate distribution functions based on pair-copulas, while Aas et al. (2009) were the first to set the PCC in an inferential context.

## PCC in three dimensions

A pair-copula construction of a three-dimensional density is given by

$$
\begin{aligned}
f\left(x_{1}, x_{2}, x_{3}\right)= & f_{1}\left(x_{1}\right) \cdot f_{2}\left(x_{2}\right) \cdot f_{3}\left(x_{3}\right) \\
\cdot & c_{12}\left(F_{1}\left(x_{1}\right), F_{2}\left(x_{2}\right)\right) \cdot c_{23}\left(F_{2}\left(x_{2}\right), F_{3}\left(x_{3}\right)\right) \\
\cdot & c_{13 ; 2}\left(F_{1 \mid 2}\left(x_{1} \mid x_{2}\right), F_{3 \mid 2}\left(x_{3} \mid x_{2}\right)\right)
\end{aligned}
$$

## Special case: Trivariate normal distribution

If the marginal distributions are standard normal and $c_{12}, c_{23}$ and $c_{13 ; 2}$ are bivariate Gaussian copula densities, the resulting distribution is trivariate normal.

## PCC in five dimensions

A possible pair-copula construction of a five-dimensional density is given by

$$
\begin{array}{ll} 
& f\left(x_{1}, x_{2}, x_{3}, x_{4}, x_{5}\right) \\
= & f\left(x_{1}\right) \cdot f\left(x_{2}\right) \cdot f\left(x_{3}\right) \cdot f\left(x_{4}\right) \cdot f\left(x_{5}\right) \\
\cdot & c_{12}\left(F\left(x_{1}\right), F\left(x_{2}\right)\right) \cdot c_{23}\left(F\left(x_{2}\right), F\left(x_{3}\right)\right) \cdot c_{34}\left(F\left(x_{3}\right), F\left(x_{4}\right)\right) \cdot c_{45}\left(F\left(x_{4}\right), F\left(x_{5}\right)\right) \\
\cdot & c_{13 ; 2}\left(F\left(x_{1} \mid x_{2}\right), F\left(x_{3} \mid x_{2}\right)\right) \cdot c_{24 ; 3}\left(F\left(x_{2} \mid x_{3}\right), F\left(x_{4} \mid x_{3}\right)\right) \cdot c_{35 ; 4}\left(F\left(x_{3} \mid x_{4}\right), F\left(x_{5} \mid x_{4}\right)\right) \\
\cdot & c_{14 ; 23}\left(F\left(x_{1} \mid x_{2}, x_{3}\right), F\left(x_{4} \mid x_{2}, x_{3}\right)\right) \cdot c_{25 ; 34}\left(F\left(x_{2} \mid x_{3}, x_{4}\right), F\left(x_{5} \mid x_{3}, x_{4}\right)\right) \\
\cdot & c_{15 ; 234}\left(F\left(x_{1} \mid x_{2}, x_{4}, x_{3}\right), F\left(x_{5} \mid x_{2}, x_{4}, x_{3}\right)\right) .
\end{array}
$$

There are as many as 480 different such constructions in the five-dimensional case, 23,040 in the 6 -dimensional case and $2,580,480$ in the 7 -dimensional case.

## Regular vines

- Hence, for high-dimensional distributions, there are a significant number of possible pair-copula constructions.
- To help organising them, Bedford and Cooke (2001) introduced graphical models denoted regular vines (R-vines).


## Regular vine (Bedford and Cooke 2002)

A regular vine is a sequence of $d-1$ linked trees where:

- Tree $T_{1}$ is a tree on nodes 1 to $d$.
- Tree $T_{j}$ has $d+1-j$ nodes and $d-j$ edges.
- Edges in tree $T_{j}$ become nodes in tree $T_{j+1}$.
- Proximity condition: Two nodes in tree $T_{j+1}$ can be joined by an edge only if the corresponding edges in tree $T_{j}$ share a node.


## Example in five dimensions



Density

$$
\begin{aligned}
f= & f_{1} \cdot f_{2} \cdot f_{3} \cdot f_{4} \\
& \cdot c_{14} \cdot c_{15} \cdot c_{24} \cdot c_{34} \\
& \cdot c_{12 ; 4} \cdot c_{13 ; 4} \cdot c_{45 ; 1} \\
& \cdot c_{23 ; 14} \cdot c_{35 ; 14} \\
& \cdot c_{25 ; 134}
\end{aligned}
$$



## Matrix representation



## Matrix

Morales-Napoles (2008) shows how a lower triangular matrix may be used to store a regular vine.

$$
M=\left(\begin{array}{lllll}
5 & & & & \\
2 & 2 & & & \\
3 & 3 & 3 & & \\
4 & 1 & 1 & 1 & \\
1 & 4 & 4 & 4 & 3
\end{array}\right)
$$

$T_{4}$

## Special cases: C and D-vines

C-vine: Each tree has a unique node connected to $d-j$ edges.

$$
\begin{aligned}
f_{1234} & =f_{1} \cdot f_{2} \cdot f_{3} \cdot f_{4} \\
& \cdot c_{12} \cdot c_{13} \cdot c_{14} \\
& \cdot c_{23 ; 1} \cdot c_{24 ; 1} \\
& c_{34 ; 12}
\end{aligned}
$$

Useful for ordering by importance

D-vine: No node is connected to more than 2 edges.

$$
\begin{aligned}
f_{1234} & =f_{1} \cdot f_{2} \cdot f_{3} \cdot f_{4} \\
& \cdot c_{12} \cdot c_{23} \cdot c_{34} \\
& \cdot c_{13 ; 2} \cdot c_{24 ; 3} \\
& \cdot c_{14 ; 23}
\end{aligned}
$$

Useful for temporal ordering of variables

tree 1
tree 2



## General density expressions

- C-vine (Aas et al. 2009)

$$
f\left(x_{1}, \ldots x_{d}\right)=\left[\prod_{k=1}^{d} f\left(x_{k}\right)\right] \times\left[\prod_{j=1}^{d-1} \prod_{i=1}^{d-j} c_{j, j+i ; 1, \ldots, j-1}\right]
$$

- D-vine (Aas et al. 2009)

$$
f\left(x_{1}, \ldots x_{d}\right)=\left[\prod_{k=1}^{d} f\left(x_{k}\right)\right] \times\left[\prod_{j=1}^{d-1} \prod_{i=1}^{d-j} c_{i, i+j ; i+1, \ldots, i+j-1}\right]
$$

- Regular vine (Dißmann et al. 2013)

$$
f\left(x_{1}, \ldots, x_{d}\right)=\left[\prod_{k=1}^{d} f_{k}\left(x_{k}\right)\right] \times\left[\prod_{j=d-1}^{1} \prod_{i=d}^{j+1} c_{m_{j, j}, m_{i, j} ; m_{i+1, j}, \ldots, m_{n, j}}\right]
$$

Here, $m_{i, j}$ refers to element $(i, j)$ in the matrix representation of the R -vine.

## Conditional distribution functions

- The conditional distributions needed as copula arguments at level $j$ are obtained as partial derivatives of the copulae at level $j-1$.
- This is due to the following result from Joe (1996) stating that under regularity conditions, we have

$$
F(x \mid \mathbf{v})=\frac{\partial C_{x v_{j} ; \mathbf{v}_{-j}}\left(F\left(x \mid \mathbf{v}_{-j}\right), F\left(v_{j} \mid \mathbf{v}_{-j}\right)\right)}{\partial F\left(v_{j} \mid \mathbf{v}_{-j}\right)}
$$

## Building blocs

- The resulting multivariate distribution will be valid even if the bivariate copulae involved in the pair-copula construction are of different type.
- One may for instance combine the following types of pair-copulae
- Gaussian (no tail dependence)
- Clayton (lower tail dependence)
- Gumbel (upper tail dependence)
- Student (upper and lower tail dependence)






## How can we estimate and model select PCCs ?

Three problems: (Czado et al. (2013))
(1) How to estimate the pair copula parameters for a given vine tree structure and the pair copula families for each edge?
(2) How to select the pair copula families and estimate the corresponding parameters for a given vine tree structure?
(3) How to select and estimate all components of a regular vine?


## Problem 1: Parameter estimation for given tree structure and copula families

- Sequential estimation:
- Parameters are sequentially estimated starting from the top tree until the last (Aas et al. (2009), Czado et al. (2012)).
- Asymptotic theory available (Hobæk Haff (2012), Hobæk Haff (2013)), however standard error estimates are difficult to compute.
- Can be used as starting values for maximum likelihood.
- Maximum likelihood estimation:
- Asymptotically efficient under regularity conditions, estimated standard errors numerically challenging (Stoeber and Schepsmeier (2012))
- Uncertainty in value-at-risk (high quantiles) is difficult to assess.
- Bayesian estimation:
- Posterior is tractable using Markov Chain Monte Carlo (Min and Czado (2011) for D-vines and Gruber (2011) for R-vines)
- Prior beliefs can be incorporated and credible intervals allow to assess uncertainty for all quantities.


## How does sequential and ML estimation work ?

Parameters: $\Theta=\left(\theta_{12}, \theta_{23}, \theta_{13 ; 2}\right)$
Copula observations: $\left\{\left(u_{1 t}, u_{2 t}, u_{3 t}\right), t=1, \cdots, T\right\}$

## Sequential estimates:

- Estimate $\theta_{12}$ from $\left\{\left(u_{1 t}, u_{2 t}\right), t=1, \cdots, T\right\}$
- Estimate $\theta_{23}$ from $\left\{\left(u_{2 t}, u_{3 t}\right), t=1, \cdots, T\right\}$.
- Define pseudo observations

$$
\hat{u}_{1 \mid 2 t}:=F\left(u_{1 t} \mid u_{2 t}, \hat{\theta}_{12}\right) \text { and } \hat{u}_{3 \mid 2 t}:=F\left(u_{2 t} \mid u_{3 t}, \hat{\theta}_{23}\right)
$$

Finally estimate $\theta_{13 ; 2}$ from $\left\{\left(\hat{u}_{1 \mid 2 t}, \hat{u}_{3 \mid 2 t}\right), t=1, \cdots, T\right\}$.

## Maximum likelihood

$$
\begin{aligned}
L(\Theta \mid x) & =\sum_{t=1}^{T}\left[\log c_{12}\left(u_{1 t}, u_{2 t} \mid \theta_{12}\right)+\log c_{23}\left(u_{2 t}, u_{3 t} \mid \theta_{23}\right)\right. \\
& \left.+\log c_{13 ; 2}\left(F\left(u_{1 t} \mid u_{2 t}, \theta_{12}\right), F\left(u_{2 t} \mid u_{3 t}, \theta_{23}\right) \mid \theta_{13 ; 2}\right)\right]
\end{aligned}
$$

## Problem 2: Joint estimation of pair copula families and parameters

- Classical approach:
- Restrict to a set of bivariate pair copula families and use AIC or Vuong test to select family
- Check for truncation possibilities (Brechmann et al. (2012)) by using independence copulas in higher trees
- Bayesian approach:
- Reversible jump (RJ) MCMC (Min and Czado (2011))
- MCMC with model indicators (Smith et al. (2010)) choosing between an independence copula and a fixed copula family.

Only one more problem to go ...


> sequential treewise approach
> (see Dißmann et al. (2013))

## How does this treewise selection of R-vines work?

Idea: Capture strong pairwise dependencies first
For Tree $\ell=1, \ldots, d-1$
(1) Calculate an empirical dependence measure $\hat{\delta}_{j k \mid D}$ for all variable pairs $\{j k \mid D\}$ ( $\rightarrow$ edge weights: Kendall's $\tau$, tail dependence coefficients) allowed by the proximity condition ( $D$ is empty for Tree 1 ).
(2) Select the tree on all nodes that maximizes the sum of absolute empirical dependencies ( $\rightarrow$ maximum spanning tree)
Choose independence copula if possible.
(3) For each selected edge $\{j, k\}(\{j, k\} \mid D)$ in Tree 1 (in Tree $\ell>1$ ), select a copula and estimate the corresponding parameter(s).
(9) Then transform to pseudo observations $F_{j \mid k \cup D}\left(u_{i j} \mid \mathbf{u}_{i, k \cup D}, \hat{\theta}_{j, k ; D}\right)$ and $F_{k \mid j \cup D}\left(u_{i k} \mid \mathbf{u}_{i, j \cup D}, \hat{\theta}_{j, k ; D}\right), i=1, \ldots, n$.

## How does this look like for Tree 1?

(1) Pairwise dependencies.

(2) Maximum dependence tree.


Czado, Jeske, and Hofmann (2013) compare sequential selection strategies

## Sequential Bayesian model selection of regular vine copulas (Gruber and Czado 2013)

- Tree by tree selection to reduce search space
- Reversible jump MCMC to select tree, pair copulas and parameters jointly
- dynamic barrier payouts based on basket of 9 Dow Jones stocks

solid: bootstrapped observed, dashed: regular vine, dotted: t copula


## Market risk model for largest Norwegian bank, DNB:

- 19 financial variables that constitute the market portfolio of DNB.
- Daily log returns from March 2003 to March 2008 (1107 obs.) are used.

| ID | description | ID | description |
| :--- | :--- | :--- | :--- |
| V1 | Norwegian Financial Index | V12 | 5-year US Government Rate |
| V2 | USD-NOK exchange rate | V13 | Norwegian bond index (BRIX) |
| V3 | EURO-NOK exchange rate | V14 | Citigroup World Government |
| V4 | YEN-NOK exchange rate |  | Bond Index (WGBI) |
| V5 | GBP-NOK exchange rate | V15 | Norwegian 6-year Swap Rate |
| V7 | 3-month Norwegian Inter | V16 | ST2X - Government Bond Index |
|  | Bank Offered Rate |  | (fix modified duration of 0.5 years) |
| V8 | Norwegian 5-year Swap Rate | V17 | Morgan Stanley World Index (MSCI) |
| V9 | 3-month Euro Interbank | V18 | OSEBX - Oslo Stock Exchange |
|  | Offered Rate |  | main index |
| V10 | 5-year German Government Rate | V19 | Oslo Stock Exchange Real Estate Index |
| V11 | 3-month US Libor Rate | V20 | S\&P Hedge Fund Index |

## Modelling procedure :

- Fit appropriate ARMA-GARCH models for log-return time series.
- Fit an R-vine as well as a multivariate Student-t copula for comparison to standardized residuals
- Pair-copulas are selected from a range of 11 bivariate families using AIC: Independence copula, Gaussian, t, Clayton, rotated Clayton (90), Gumbel, rotated Gumbel (90), Frank, Joe, Clayton-Gumbel (BB1), Joe-Clayton (BB7).


## First tree of R-vine:



## Results:

| Copula | Log <br> likelihood | No. of <br> param. | AIC |
| :---: | :---: | :---: | :---: |
| R-vine | 6390.75 | 92 | -12597.50 |
| Student-t | 6324.98 | 172 | -12305.96 |

## Number of parameters:

Note that the number of parameters to be estimated for a 19-dimensional R -vine usually is at least $\mathrm{d}(\mathrm{d}-1) / 2$. The reason why the number in the table is 92 and not 171 is that a large amount of the pair-copulae in this R -vine are identified as the independence copula, using the bivariate independence test based on Kendall's tau as described in Genest and Favre (2007) .

## Truncation(1):

- The number of parameters in an R -vine grows quadratically with the dimension.
- Hence, it would be useful to be able to reduce the model complexity.
- In Brechmann et al. (2012) we have studied the problem of determining whether an R -vine may be truncated.
- By a truncated R -vine at level $K$, we mean an R -vine with all pair-copulae with conditioning set larger than or equal to $K$ set to independence copulae.
- We fit one tree at a time and use the likelihood ratio test of Vuong (1989) to determine whether an additional tree provides a significant gain in the model fit.


## Results:

| Copula | Log <br> likelihood | No. of <br> param. | AIC |
| :--- | ---: | ---: | ---: |
| R-vine | 6390.75 | 92 | -12597.50 |
| Student-t | 6324.98 | 172 | -12305.96 |
| 6-level R-vine | 6274.47 | 77 | -12394.94 |
| 4-level R-vine | 6234.05 | 68 | -12332.10 |

## Conclusion:

We conclude from this that the most important dependencies in this data set are actually captured in the first four to six trees, meaning that the corresponding R-vine copula may be truncated at level 6 , or even at level 4, depending on the desired level of parsimonity (and of course at the expense of accuracy).

## Recent advances for vines

- Simplified and non simplified vines
- Time varying regular vines
- Discrete and discrete/continuous vines
- Non Gaussian DAG's using pair copula constructions
- Vines with non parametric pair copulas: Haff and Segers (2013), Kauermann and Schellhase (2013)
- Acceleration of MCMC algorithms: Schmidl et al. (2013)


## Simplified and non-simplified vines

## Simplifying assumption

Pair copulas depend on their conditioning value only through their conditional distributions (Haff, Aas, and Frigessi 2010)

- Simplified vine copulas:
- multivariate Gauss copula
- multivariate t-copula only one arising from arising from scale mixture of normals (Stöber et al. 2012)
- multivariate Clayton is the only one among the Archimedean copulas (Stöber et al. 2012)
- Non-simplified vines:
- Acar et al. (2012) use a smoothing approach to deal with non simplified vines in 3 dimensions based on Acar et al. (2011).
- Occurs when considering one factor models: $X_{j}=Z_{0}+X_{j}$ for $j=1, \ldots, d$.


## Effects of simplifying assumption

Trivariate extension of FGM copula (Stöber 2013)




- left: Kulback Leibler distance
- middle: sample size needed to distinguish the models in a LR test
- right: relative difference in Value-at-Risk


## Violation of simplifying assumption might indicate time varying dependence

Conditional Kendalls $\tau$ rank correlation between the USD/EUR and USD/CAD return exchange rate conditional on USD/GBP being in a given decile of the distribution
$\tau_{\text {EUR,CAD|GBP }}$


The boxplots are obtained using a non-parametric bootstrap. The analysis in Stöber and Czado (2013) shows that data has time varying dependence.

## Time varying regular vines

- AR(1) copula dynamics
- Bayesian bivariate analysis: Almeida and Czado (2011)
- Multivariate analysis: Almeida et al. (2012)
- Regime switching
- C-vine, copula parameters only, EM: Chollete et al. (2008)
- R-vine and copula parameter switches, EM, MCMC: Stöber and Czado (2013)
- Marginal and copula switches: Stöber (2013)


## Smoothed probabilities of being in non-Gaussian regime



The solid lines correspond to Bayesian MCMC estimates, the dotted lines to $90 \%$ Cls (Stöber 2013)

## PCC based network models

- Bayesian belief networks: They were first considered by Hanea et al. (2006).
- Pair-copula Bayesian networks (PCBN)
- Bauer et al. (2012) used a PCC construction to build Non Gaussian DAG models.
- Bauer (2013) and Bauer and Czado (2012) give general algorithms to estimate PCBN and provides a PC algorithm to construct network.


Parent ordering Pair copula family Kendall's $\tau$

## Discrete and discrete/continuous vines

- Discrete vines Panagiotelis et al. (2012) construct an efficient PCC using D-vines based on the distribution function
- Discrete/continuous vines Stöber et al. (2012) and Stöber (2013) extend to cover discrete/continuous variables and allow for regular vines
- solid (diabetes, hypertension),
- dashed (diabetes, no hypertension)
- dotted (no diabetes, hypertension)
- dash-dotted (no diabetes, no hypertersion)



## Selected Applications

- Financial risk management:
- Euro Stoxx 50 (Brechmann and Czado 2012)
- Systemic risk simulation (Brechmann et al. 2013)
- Operational risk: (Brechmann et al. 2013)
- Multivariate options: (Gruber and Czado 2013)
- Realized volatility: (Vaz de Melo Mendes and Accioly 2013)
- Hydrology: (Gräler et al. 2013)
- Data mining: (Lopez-Paz et al. 2013)
- Health: comorbidity (Stöber et al. 2012)
- Environmental Science:(Gräler and Pebesma 2011) (Pachali 2011)


## What have we learned?

- Standard multivariate copulas are less flexible, while PCC's such as C-, D- and R-vines are much more flexible.
- Sequential and MLE parameter estimation of C-, D- and R-vines are available in $\mathbf{R}$ packages CDVine and VineCopula.
- Sequential and full Bayesian estimation and Bayesian model selection of vine trees and copula families for regular vines available, but need further testing and development
- Pair copula constructions can be extended to mixed continuous and discrete data.
- Vine copulas are useful for financial risk management


## What needs to be done?

- Non-parametric pair copulas, spatial vines, vines for data mining
- More applications in finance, insurance ...


## Vine resource page:

www-m4.ma.tum.de/forschung/vine-copula-models

Vine workshop book: Kurowicka and Joe (2011)

Next vine workshop: Jan. 3/4 2014, Peking, China (?)

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