Sparse topology selection in graphical models of autoregressive time series

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- Gaussian graphical models
- Graphical models of autoregressive time series
- Algorithms

n-dimensional Gaussian vector

$$x = (x_1, \dots, x_n) \sim \mathcal{N}(0, \Sigma)$$

 x_i , x_j are conditionally independent (given the rest of x) if

$$(\Sigma^{-1})_{ij} = 0$$

modeled as undirected graph with n nodes; arc i, j is absent if $(\Sigma^{-1})_{ij} = 0$



Log-likelihood function for N independent samples of x

$$\frac{N}{2} \left(\log \det \Sigma^{-1} - \mathbf{tr}(C\Sigma^{-1}) \right)$$

C is sample covariance

ML estimation of Σ , for given topology

 $\begin{array}{ll} \text{minimize} & -\log \det X + \mathbf{tr}(CX) \\ \text{subject to} & \text{given sparsity pattern of } X \end{array}$

a convex problem in $X = \Sigma^{-1}$

known as covariance selection (Dempster 1972)

Topology selection via model selection criteria

• enumerate topologies and for each topology, solve ML problem

minimize
$$-\mathcal{L}(X) = -\log \det X + \mathbf{tr}(CX)$$

subject to sparsity pattern of X

• rank ML estimates $X_{ml} = \Sigma_{ml}^{-1}$ using an information criterion

AIC =
$$-2\mathcal{L}(X_{ml}) + 2k$$
 (Akaike)
AIC_c = $-2\mathcal{L}(X_{ml}) + \frac{2Nk}{N-k-1}$ (second order Akaike)
BIC = $-2\mathcal{L}(X_{ml}) + k \log N$ (Bayes)

k is number of parameters (\propto nonzeros in X); N is sample size

this approach is only feasible for small graphs

Topology selection via 1-norm regularization

Regularized ML problem

minimize
$$-\log \det X + \operatorname{tr}(CX) + \gamma \sum_{i,j} |X_{ij}|$$

Dual problem

$$\begin{array}{ll} \mathsf{maximize} & \log \det(C+Z) \\ \mathsf{subject to} & |Z_{ij}| \leq \gamma, \quad i,j=1,\ldots,n \end{array}$$

• convex; primal or dual can be solved by first-order methods

Yuan & Lin 2007; Banerjee, El Ghaoui & d'Aspremont 2008; Friedman, Hastie & Tibshirani 2008; Lu 2008, 2009; Scheinberg & Rish 2009, . . .

• choice of γ : rank topologies on trade-off curve by AIC, AIC_c or BIC other methods: Banerjee et al. 2008, Friedman et al. 2007, Ravikumar et al. 2008; Meinshausen & Bühlmann 2006

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n-dimensional stationary Gaussian time series x(t), $t \in \mathbf{Z}$

 x_i and x_j are conditionally independent (given the rest of x(t)) if

$$\left(S(\omega)^{-1}\right)_{ij} = 0$$

 $S(\omega)$ is spectral density matrix:

$$S(\omega) = \sum_{k=-\infty}^{\infty} R_k e^{-jk\omega}, \qquad R_k = \mathbf{E} x(t+k)x(t)^T \qquad (j = \sqrt{-1})$$

Brillinger 1981, Dahlhaus 2000

Example (autoregressive model of order 4)



- coherence spectrum: $S(\omega)$ normalized to have diagonal one
- partial coherence: $S(\omega)^{-1}$ normalized to have diagonal one

$$B_0 x(t) = -\sum_{k=1}^p B_k^T x(t-k) + w(t), \qquad w(t) \sim \mathcal{N}(0, I)$$

without loss of generality, assume B_0 symmetric, positive definite

Inverse spectrum

$$S(\omega)^{-1} = Y_0 + \sum_{k=1}^{p} \left(Y_k e^{-jk\omega} + Y_k^T e^{jk\omega} \right), \qquad Y_k = \sum_{l=0}^{p-k} B_l B_{l+k}^T$$

Conditional independence relations

$$(S(\omega)^{-1})_{ij} = 0 \qquad \iff \qquad (Y_k)_{ij} = (Y_k)_{ji} = 0, \quad k = 0, 1, \dots, p$$

minimize
$$-\log \det X_{00} + \operatorname{tr}(CX)$$

subject to $X = \begin{bmatrix} X_{00} & \cdots & X_{0p} \\ \vdots & \ddots & \vdots \\ X_{p0} & \cdots & X_{pp} \end{bmatrix} = \begin{bmatrix} B_0 \\ \vdots \\ B_p \end{bmatrix} \begin{bmatrix} B_0 \\ \vdots \\ B_p \end{bmatrix}^T$
given sparsity pattern of $\sum_l X_{l,l+k}$ for $k = 0, \dots, p$

- maximizes conditional likelihood (conditioned on first *p* values)
- C is sample covariance estimate from observations of x(t) (see later)
- variables are X, B_0 , . . . , B_p
- equality constraints $X_{ij} = B_i B_j^T$ make problem nonconvex

$$\begin{array}{ll} \text{minimize} & -\log \det X_{00} + \mathbf{tr}(CX) \\ \text{subject to} & X = \left[\begin{array}{ccc} X_{00} & \cdots & X_{0p} \\ \vdots & \ddots & \vdots \\ X_{p0} & \cdots & X_{pp} \end{array} \right] \succeq 0 \\ \\ \text{given sparsity pattern of } \sum_{l} X_{l,l+k} \text{ for } k = 0, \dots, p \end{array}$$

• exact if optimal X has rank n, *i.e.*, can be factored as $X_{ij} = B_i B_j^T$

- from duality: relaxation is exact if C is pos. definite and **block-Toeplitz**
- in practice: often exact even for non-Toeplitz sample covariances ${\cal C}$

minimize
$$-\log \det X_{00} + \operatorname{tr}(CX)$$

subject to $X \succeq 0$, $\mathsf{P}\left(\sum_{l} X_{l,l+k}\right) = 0$

- X is $(p+1) \times (p+1)$ symmetric block matrix with blocks X_{ij} of order n
- $\mathsf{P}(U)$ is projection on zero pattern

Dual problem

maximize
$$\log \det W + n$$

subject to $\begin{bmatrix} W & 0 \\ 0 & 0 \end{bmatrix} \preceq C + \mathsf{T}(\mathsf{P}(Z_0), \dots, \mathsf{P}(Z_p))$

- variables W, Z_0 , . . . , Z_p are $n \times n$, with W and Z_0 symmetric
- $\mathsf{T}(U_0,\ldots,U_p)$ is block-Toepliz matrix with first row U_0,\ldots,U_p

Property of block-Toeplitz matrices: if $W \succ 0$ and

$$\mathsf{T}(U_0, \dots, U_p) = \begin{bmatrix} U_0 & U_1 & \cdots & U_p \\ U_1^T & U_0 & \cdots & U_{p-1} \\ \vdots & \vdots & \ddots & \vdots \\ U_p^T & U_{p-1}^T & \cdots & U_0 \end{bmatrix} \succeq \begin{bmatrix} W & 0 & \cdots & 0 \\ 0 & 0 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & 0 \end{bmatrix}$$

then $\mathsf{T}(U_0,\ldots,U_p)\succ 0$

Complementary slackness for ML problem ($Z = T(P(Z_0), ..., P(Z_p))$)

$$W = X_{00}^{-1}, \qquad \operatorname{tr}\left(X\left(C + Z - \left[\begin{array}{cc}W & 0\\0 & 0\end{array}\right]\right)\right) = 0$$

hence, if C is block-Toeplitz, optimal X has rank n

from measurements $\hat{x}(1), \ldots, \hat{x}(N)$, estimate $R_k = \mathbf{E} x(t) x(t+k)^T$

$$C = \frac{1}{M} \sum_{t=t_1}^{t_2} \begin{bmatrix} \hat{x}(t) \\ \hat{x}(t-1) \\ \vdots \\ \hat{x}(t-p) \end{bmatrix} \begin{bmatrix} \hat{x}(t) \\ \hat{x}(t-1) \\ \vdots \\ \hat{x}(t-p) \end{bmatrix}^T \approx \begin{bmatrix} R_0 & R_1 & \cdots & R_p \\ R_1^T & R_0 & \cdots & R_{p-1} \\ \vdots & \vdots & \ddots & \vdots \\ R_p^T & R_{p-1}^T & \cdots & R_0 \end{bmatrix}$$

Non-windowed estimate: $t_1 = p + 1$, $t_2 = N$, M = N - p

- arises in conditional ML/LS estimation
- generally not block-Toeplitz but approaches block-Toeplitz for large N

Windowed estimate: $t_1 = 1$, $t_2 = N + p$, M = N

- assume $\hat{x}(t) = 0$ for t < 1 and t > N
- block-Toeplitz



- generate $50\ {\rm trials}$ of C for each model
- relaxation for (non-Toeplitz) non-windowed sample covariance matrix C

hourly values of CO, NO, NO₂, O₃, solar radiation at Azusa (N = 8370)



blue: empirical spectrum; red: optimal AR model for BIC (p = 4)

Topology selection via nonsmooth regularization

Estimation with known topology

 $\begin{array}{ll} \mbox{minimize} & -\log \det X_{00} + {\bf tr}(CX) \\ \mbox{subject to} & X \succeq 0 \\ \mbox{given sparsity pattern of } Y_k = \sum_{l=0}^{p-k} X_{l,l+k}, \quad k=0,\ldots,p \end{array}$

Nonsmooth regularization

minimize
$$-\log \det X_{00} + \operatorname{tr}(CX) + \gamma h(Y_0, \dots, Y_p)$$

subject to $X \succeq 0$

convex penalty h promotes common, symmetric sparsity pattern of Y_k :

$$h(Y_0, Y_1, \dots, Y_p) = \sum_{i>j} \max_k \max\{|(Y_k)_{ij}|, |(Y_k)_{ji}|\}$$

Example

 $n=20\text{, }p=2\text{, exact }S(\omega)^{-1}$ has 76 nonzeros



Regularized maximum likelihood problem (512 samples)

minimize
$$-\mathcal{L}(X) + \gamma h(Y)$$
 $(\mathcal{L}(X) = \log \det X_{00} - \operatorname{tr}(CX))$



Topology for $\gamma=0.15$



- blue squares: correctly classified
- red circles: incorrectly classified as nonzero
- plus signs: incorrectly classified as zero

threshold the inverse spectrum from one of three estimation methods

- ML estimation (a.k.a. LS estimation): solve Yule-Walker equations
- ML estimation with added Tikhonov regularization term

$$\gamma \sum \|B_k\|_F^2 = \gamma \operatorname{tr} X$$

equivalent to ML estimate using $C:=C+\gamma I$

• ML estimation with added nonsmooth regularization term h(Y)

Error in topology as function of sample size



- topology estimated by applying threshold to ML estimate (LS), Tikhonov-regularized ML estimate, and *h*-regularized ML estimate (L1)
- graphs show fraction of entries misclassified as zero/nonzero

Example: 17 stock market indices



detrended daily returns of 17 market indices during 1997-1999; N = 540



link widths show $\rho_{ij} = \max |R_{ij}(\omega)|$, $R(\omega)$ is normalized inverse spectrum

Dataset (Mitchell, 2004)

- time series of length N = 640
- 1718 voxels in ROI, reduced to n = 7, 50, 100, 190 by averaging groups
- two inputs: 'picture', 'sentence'

Experiment

- estimate sparse models for the two inputs by regularized ML estimation
- validate models by an input classification experiment

Model selection from regularized ML estimation plus BIC

• model order

input	n = 7	n = 50	n = 100	n = 190
picture	p = 1	p = 1	p = 0	p = 0
sentence	p=1	p = 1	p = 0	p = 0

• sparsity



Model validation via input classification

- use selected models to classify the two inputs from unseen data
- select the input with the highest likelihood

Classification error versus model size

model order	n = 7	n = 50	n = 100	n = 190
p = 0	21%	16%	11%	6%
p=1	20%	16%	16%	11%

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$$\begin{array}{ll} \text{minimize} & -\log \det X_{00} + \operatorname{tr}(CX) + \gamma h(Y) \\ \text{subject to} & X = \begin{bmatrix} X_{00} & \cdots & X_{0p} \\ \vdots & \ddots & \vdots \\ X_{p0} & \cdots & X_{pp} \end{bmatrix} \succeq 0 \\ & Y_k = \sum_{l=0}^{p-k} X_{l,l+k}, \quad k = 0, \dots, p \end{array}$$

- variables X and $Y = (Y_0, \ldots, Y_p)$ with Y_k and X_{ij} square of order n
- $h(Y_0, Y_1, \ldots, Y_p)$ is nonsmooth penalty (sum of infinity norms)

$$h(Y) = \sum_{j>i} \max_{k} \max\{|Y_{k,ij}|, |Y_{k,ij}|\}$$

minimize
$$f(C + T(Z))$$

subject to $\sum_{k=0}^{p} (|Z_{k,ij}| + |Z_{k,ji}|) \le \gamma, \quad i \ne j$

variable $Z = (Z_0, Z_1, \ldots, Z_p)$, Z_k square with zero diagonal

- constraints are independent 1-norm constraints
- T(Z) is block-Toeplitz matrix with first row Z

•
$$f(V) = -\log \det(V_{00} - V_{1:p,0}^T V_{1:p,1:p}^{\dagger} V_{1:p,0})$$
 if

$$V = \begin{bmatrix} V_{00} & \cdots & V_{p0}^T \\ \vdots & \ddots & \vdots \\ V_{p0} & \cdots & V_{pp} \end{bmatrix} = \begin{bmatrix} V_{00} & V_{1:p,0}^T \\ V_{1:p,0} & V_{1:p,1:p} \end{bmatrix} \succeq 0$$

• objective is convex differentiable, and **closed** if C is block-Toeplitz

 $\begin{array}{ll} \text{minimize} & g(x) \\ \text{subject to} & x \in \mathcal{C} \end{array}$

g convex, differentiable; C a 'simple' convex set (e.g., 1-norm ball)

Basic algorithm

$$x^{(k+1)} = \mathcal{P}\left(x^{(k)} - \alpha \nabla g(x^{(k)})\right)$$

 ${\mathcal P}$ is (inexpensive) projection on ${\mathcal C}$

Accelerated algorithms (Nesterov, Beck and Teboulle, Tseng, . . .)

- same complexity per step; faster ('optimal') convergence in theory
- known convergence theory does not apply to our problem

'Arc search': backtracking search for α



initialize α by Barzilai-Borwein stepsize

(Straight) line search

$$x := (1 - \alpha)x + \alpha \mathcal{P} \left(x - t \nabla g(x) \right)$$

 α determined by backtracking

Numerical example



- example with n = 300, p = 2 (225150 variables)
- exact FISTA uses decreasing stepsize (required in convergence proof)
- modified FISTA is FISTA with non-monotone stepsize

Estimation with known topology

- convex relaxation of constrained ML estimation problem
- relaxation is exact if sample covariance matrix C is block-Toeplitz
- in practice, it is often exact for almost-Toeplitz ${\cal C}$

Topology selection

- convex nonsmooth (ℓ_1 -type) regularization of ML problem
- useful heuristic for reducing #models ranked by information criteria
- efficient solution via first-order methods applied to the dual

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