Conditional sampling for max-stable random fields¹

Summary

Max-stable random fields play a central role in modeling extreme value phenomena. We obtain an explicit formula for the conditional probability in general max-linear models, which include a large class of max-stable random fields. As a consequence, we develop an algorithm for efficient and exact sampling from the conditional distributions. Our method provides a computational solution to the prediction problem for spectrally discrete max-stable random fields. This work offers new tools and a new perspective to many statistical inference problems for spatial extremes, arising, for example, in meteorology, geology, and environmental applications.

Introduction

A max-stable random field with α -Fréchet marginals has the representation

$$X_t = \int_U^{\mathbf{e}} f_t(u) M_\alpha(du), \quad f_t \in L^\alpha_+(U,\mu)$$

with $t \in T$, where T can be \mathbb{R} , \mathbb{R}^2 , \mathbb{R}^d , etc. (see e.g. [4] & [7]). The finite-dimensional distributions of $X = \{X_t\}_{t \in T}$ are (for $t_i \in T, x_i > 0$) as follows:

$$\mathbb{P}\{X_{t_i} \le x_i, \ 1 \le i \le n\} = \exp\left\{-\int_U \max_{1 \le i \le n} \left(\frac{f_{t_i}(u)}{x_i}\right)^{\alpha} \mu(du)\right\}$$
$$= \exp\left\{-\int_{\mathbb{S}^{n-1}_+} \left(\max_{1 \le i \le n} w_i/x_i\right)^{\alpha} \Gamma(dw)\right\}.$$

Here $\mathbb{S}^{n-1}_+ = \{ w \in \mathbb{R}^n_+ : \max_{1 \le i \le n} w_i = 1 \}$ is the positive unit sphere in the sup-norm, and Γ is a unique finite measure on \mathbb{S}^{n-1}_{+} called the *spectral measure* of the distribution (see e.g. [5]).

The Max–Linear Model

Let $a_{i,j} \ge 0$ and Z_j , $1 \le j \le p$ be i.i.d. α -Fréchet.

$$X_i := \bigvee_{j=1}^p a_{i,j} Z_j, \quad (1 \le i \le n)$$

denoted by

 $X = A \odot Z$, with $X = (X_i)_{i=1}^n$, $Z = (Z_j)_{j=1}^p$, $A = (a_{i,j})_{j=1}^n$

is called a max-linear model.

Example Let $\phi_t(j)$, $t \in T$, $1 \leq j \leq p$ be non-negative functions. Consider the max-stable $(\alpha - Féchet)$ random field:

$$X_t := \bigvee_{j=1}^p \phi_t(j) Z_j, \quad t \in T.$$

The random vectors $(X_{t_i})_{i=1}^n$ follow the max-linear model in (3) with $A = (\phi_{t_i}(j))_{n \times p}$. Remarks

- Since the Z_i 's are i.i.d. α -Fréchet, the max-linear model has the representation (1) with simple spectral functions f_t 's.
- The spectral measure Γ for the max-linear model is *discrete*. Hence the models (3) and (4) are spectrally discrete.
- The max-linear model in (3) can approximate (in \mathbb{P}) any general model (1).

Prediction for Max–Stable Random fields

The Problem: A max-stable random field $\{X_t\}_{t\in T}$ is observed at locations $t_1, \dots, t_n \in T$. The goal is to predict, X_{s_1}, \cdots, X_{s_m} .

Remarks

- Previously addressed by [3, 1, 2] among others.
- Important but very challenging problem.
- No hope for an *exact* analytical solution.

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A Computational Solution

In our recent work [8], we obtain a computational solution of the general prediction problem for the *max–linear* model.

- We focus on the spectrally discrete max-linear model (3).
- Obtain an exact formula (see [8]) of the regular conditional

$$Z|X = x$$

• Thus, obtain the conditional distribution:

(6)

Y|X = x, where $Y = B \odot Z$, with $Y = (Y_k)_{k=1}^m$, $B = (b_{k,j})_{m \times p}$, where B is a given non–negative matrix.

- Sample efficiently and exactly (with a computer) from (5) and hence (6).
- Thus, obtain empirical (Monte Carlo) approximations for functionals of Y|X = x.
- Can obtain L^1, L^2 , etc. optimal predictors such as conditional medians, means, etc.; conditional quantiles and confidence sets (not just intervals!).

An algorithm for efficient and exact sampling from the conditional distribution is available in C/C++ with interface as an R-package from Yizao Wang's web page:

http://www.stat.lsa.umich.edu/~yizwang/software/maxLinear/

Prediction for MARMA

The max-autoregressive moving averages (MARMA(m,q)) are the stationary time series solutions to: $X_t = \phi_1 X_{t-1} \vee \cdots \vee \phi_m X_{t-m} \vee Z_t \vee \theta_1 Z_{t-1} \vee \cdots \vee \theta_q Z_{t-q}, \quad t \in \mathbb{Z},$ with $\phi_i \ge 0, \theta_j \ge 0, 1 \le i \le m, 1 \le j \le q$ and i.i.d. 1-Fréchet Z_t 's. If $\phi^* = \bigvee_{i=1}^m \phi_i < 1$, then (7) has a unique solution:

$$X_t = \bigvee_{j=0}^{\infty} \psi_j Z_{t-j}, \quad \text{with } \psi_j = \bigvee_{k=0}^{j \wedge q} \alpha_{j-k} \theta_k$$

where $\alpha_j = 0$, j < 0, $\alpha_0 = 1$ and $\alpha_j = \phi_1 \alpha_{j-1} \lor \phi_2 \alpha_{j-2} \lor \cdots \lor \phi_m \alpha_{j-m}$, $j \ge 1$. Since the ψ_i 's decay exponentially, by truncation, one can approximate well

$$X_t \stackrel{\mathbb{P}}{\approx} \widetilde{X}_t = \bigvee_{j=0}^M \psi_j Z_{t-j}$$

by a max-linear model.

The graph illustrates prediction of a MARMA(3,0) with $\phi_1 = 0.7, \phi_2 = 0.5$ and $\phi_3 = 0.3$. Observed are the first 100 values and predicted are the next 50.



The table shows the coverage and width of 95% upper prediction intervals for MARMA. The third column reports

$$\mathbb{P}\{X_{100+s} \le \widehat{X}_{100+s}^{(DR)} | X_t, \ 1 \le t$$

where $\widehat{X}_{100+s}^{(DR)}$ are the Projection Predictors of Davis & Resnick [3]. Remarks

- The accurate coverages demonstrate the validity of our algorithm.
- Naturally, the width of the prediction intervals grows with the lag.
- One can numerically quantify the prediction error of previous projection predictors of [3].

(1)

(2)

$$_{j})_{n \times p} \tag{3}$$





probability distribution of	
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(7)

Proj Pred
0.706
0.503
0.356
0.253
0.178
0.029
0.001
0.000

0.000

 $\leq 100\},$

The Discrete Smith Model

Take in (1) $T = \mathbb{R}^2$, $U = \mathbb{R}^2$ and $f_t(u) := \phi(t - u)$ with

$$\phi(t_1, t_2) := \frac{\beta_1 \beta_2}{2\pi \sqrt{1 - \rho^2}} \exp\left\{-\frac{\beta_2 \beta_2}{2\pi \sqrt{1 - \rho^2}} \exp\left\{-\frac{\beta_2 \beta_2}{2\pi \sqrt{1 - \rho^2}}\right\}\right\} + \frac{\beta_1 \beta_2}{2\pi \sqrt{1 - \rho^2}} \exp\left\{-\frac{\beta_2 \beta_2}{2\pi \sqrt{1 - \rho^2}} \exp\left\{-\frac{\beta_2 \beta_2}{2\pi \sqrt{1 - \rho^2}}\right\}\right\}$$



Rainfall Maxima in Bourgogne²



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 $-\frac{1}{2(1-\rho^2)} \Big\{ \beta_1^2 t_1^2 - 2\rho \beta_1 \beta_2 t_1 t_2 + \beta_2^2 t_2^2 \Big\} \Big\}.$

Smith moving maxima random field model

- Parameters: $\beta_1, \beta_2 > 0$ and $-1 < \rho < 1$. • To apply our conditional sampling method-
- ology, we discretize the extremal integral:

 $X_t := \bigvee h^{2/\alpha} \phi(t - u_{j_1 j_2}) Z_{j_1 j_2},$ $-q \leq j_1, j_2 \leq q-1$

where $u_{j_1,j_2} = ((j_1 + 1/2)h, (j_2 + 1/2)h).$ • Figure on the left: 4 conditional samples with $\beta_1 = \beta_2 = 1$, $\rho = 0$, given 7 observed values (all equal to 5).

- 100

- The discrete Smith model is applied to 50– year maxima at 146 stations in Bourgogne, France.
- Data marginally transformed to a (nonstandard) Gumbel scale.
- The heatmap shows the conditional median based on 5 observations with white crosses.
- The shades of the circles indicate true values.
- Crossed circles are outside a two-sided 95% prediction interval.
- The Smith model was fit using crossvalidation

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